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# Machine learning and the cross-section of emerging market stock returns<sup>☆</sup>

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## ABSTRACT

This paper compares various machine learning models to predict the cross-section of emerging market stock returns. We document that allowing for non-linearities and interactions leads to economically and statistically superior out-of-sample returns compared to traditional linear models. Although we find that both linear and machine learning models show higher predictability for stocks associated with higher limits to arbitrage, we also show that this effect is less pronounced for non-linear models. Furthermore, significant net returns can be achieved when accounting for transaction costs, short-selling constraints, and limiting our investment universe to big stocks only.

## 1. Introduction

Machine learning algorithms have been available for a long time. However, due to increased computing power and data availability, decreased data storage costs, and algorithmic innovations in recent years (cf., [Rasekhschaffe and Jones, 2019](#)), machine

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learning methods see increasing popularity in research fields such as economics, finance, and accounting.<sup>1</sup>

This paper compares various machine learning models to predict the cross-section of emerging market stock returns. More specifically, we analyze the predictive power of nine algorithms: ordinary least squares regression and elastic net as examples for traditional linear models; tree-based models such as gradient-boosted regression trees and random forest; and neural networks with one to five layers. Furthermore, we investigate the performance of an ensemble comprising the five different neural networks and an ensemble of methods that allow for non-linearities and interactions, i.e., the two tree-based models and the ensemble of neural networks. In the remainder of the paper, we often use the term ‘machine learning’ only for the two tree-based methods, the neural networks, and the two ensembles. Our data set contains stocks from 32 emerging market countries and the 36 firm-level characteristics from Kelly et al. (2019) and Windmüller (2022) falling into categories such as value, past returns, investment, profitability, intangibles, and trading frictions. The data sample covers the sample period from July 1995 to December 2021, while our 20-year out-of-sample period is from January 2002 to December 2021.

Our main findings can be summarized as follows. First, we document that the different prediction algorithms pick up similar characteristics. However, we observe that tree-based methods and neural networks also identify non-linearities and interactions of characteristics. In contrast, linear methods are restricted to linear relationships and do not allow for interactions among characteristics.

Second, return forecasts based on machine learning models lead to economically and statistically superior out-of-sample long-short returns compared to traditional linear models. Furthermore, the Fama and French (2018) six-factor model can only partly explain these long-short returns, and their alphas remain highly significant. These findings are robust to several methodological choices and for emerging market subregions. Finally, we document that machine learning forecasts beat linear models consistently over our sample period, and we cannot observe a decline in predictability over time.

Third, developed market long-short returns based on machine learning forecasts derived in the same way as their emerging market counterparts cannot explain emerging market out-of-sample returns. However, models estimated solely on developed markets data also predict emerging market stock returns. These findings indicate that similar relationships between firm characteristics and future stock returns exist for developed and emerging markets but that the pricing of these characteristics is not fully integrated between developed and emerging markets.

Fourth, the high returns of the machine learning strategies in emerging markets do not primarily stem from higher-risk months and do not revert quickly, suggesting that an underreaction explanation is more likely than a risk-based explanation. Furthermore, both linear and machine learning models show higher predictability for stocks associated with higher limits to arbitrage. However, we also document that this effect is less pronounced for machine learning forecasts than for linear regression forecasts, indicating that the superiority of machine learning models in emerging markets does not stem from limits to arbitrage.

Finally, accounting for transaction costs, short-selling constraints, and limiting our investment universe to big stocks only, we document that machine learning-based return forecasts can lead to significant net outperformance over the market and net alphas, at least when efficient trading rules are applied.

This paper contributes to the literature in at least three aspects. First, we contribute to the rapidly expanding literature on predicting the cross-section of stock returns with machine learning methods. Rasekhschaffe and Jones (2019), Freyberger et al. (2020), and Gu et al. (2020) document that more complex machine learning models are superior to linear models for the U.S. Tobek and Hronec (2020) and Drobetz and Otto (2021) find similar evidence for developed markets and Europe, respectively. However, none of the studies mentioned above investigates emerging markets. Emerging markets are important as they account for around 58% of the global gross domestic product (GDP), which is forecasted to rise to 61% by 2026.<sup>2</sup> Furthermore, the same risk factors should apply to these markets under the hypothesis that developed markets are integrated. Therefore, similar results within developed markets are not surprising, and emerging markets provide an attractive alternative for out-of-sample tests in terms of independent and new samples.

Two contemporaneously written papers, Azevedo et al. (2022) and Cakici et al. (2022a), also include emerging markets in their analysis next to developed markets. While Azevedo et al. (2022) also find that most machine learning models outperform a linear combination of anomalies, their results do not discriminate between emerging and other markets. Therefore, their results are mainly driven by developed markets. In contrast to our study, Cakici et al. (2022a) do not find superior forecasts for machine learning models compared to linear models. A potential reason for this difference might be that they train their models for each country separately while we train our models on a pooled sample of countries. However, more data might be necessary for more complex models to robustly identify non-linearities and interactions in the data.<sup>3</sup> We provide some supportive evidence for this claim by documenting that

<sup>1</sup> For instance, machine learning methods are applied to predict stock returns in Moritz and Zimmermann (2016), Rasekhschaffe and Jones (2019), Freyberger et al. (2020), Gu et al. (2020), Tobek and Hronec (2020), Chen et al. (2023), Drobetz and Otto (2021), Leippold et al. (2022), Azevedo et al. (2022), Cakici et al. (2022a), and Rubesam (2022), stock market betas in Drobetz et al. (2021), country stock returns in Cakici and Zaremba (2022), industry stock returns in Rapach et al. (2019), option returns in Bali et al. (2023), corporate bond returns in Kaufmann et al. (2021) and Bali et al. (2022), the equity premium in Rossi (2018), Treasury bond returns in Bianchi et al. (2021b) and Bianchi et al. (2021a), commodity returns in Struck and Cheng (2020), short-term bitcoin returns in Jaquart et al. (2021), cryptocurrency returns in Cakici et al. (2022b), (changes) in future company profitability in Anand et al. (2019), Van Binsbergen et al. (2020) and Chen et al. (2022), peer-implied market capitalizations in Hanauer et al. (2022a), mutual fund selection in Kaniel et al. (2022), hedge fund selection in Wu et al. (2021), mortgage risk in Sadhwani et al. (2021), or corporate directors in Erel et al. (2021).

<sup>2</sup> See, IMF, World Economic Outlook database, April 2022, <https://www.imf.org/en/Publications/WEO/weo-database/2022/April>.

<sup>3</sup> While a linear model asks for a single parameter for each predictor, in the case of non-linear models, the number of parameters to estimate rapidly expands even with a moderate number of predictors (cf., Gu et al., 2020; Hanauer et al., 2022a). As such, pooling data across countries will arguably improve the observations-to-parameters ratio.

models trained on emerging market subregions underperform models trained on the pooled sample of emerging market subregions and that the performance loss is more pronounced for machine learning models and smaller subregions. Finally, [Leippold et al. \(2022\)](#) show that machine learning models dominate linear models for Chinese A-shares. In contrast, our sample purposely excludes Chinese A-shares to represent an international investor's investable emerging market universe: for the majority of our sample period, the China A-share market was only accessible to local investors and only gradually opened up to international investors (cf., [Jansen et al., 2021](#)).

Second, we add to the literature on the drivers of emerging market stock returns. [Bekaert and Harvey \(1995\)](#) and [Harvey \(1995\)](#) were among the first to investigate emerging market country returns and their market integration. Early studies on the cross-section of emerging market stocks, such as [Rouwenhorst \(1999\)](#), [van der Hart et al. \(2003\)](#), [van der Hart et al. \(2005\)](#), [Griffin et al. \(2010\)](#), [Cakici et al. \(2013\)](#), [Hanauer and Linhart \(2015\)](#) mainly focus on size, value, and momentum. Later studies such as [Zaremba and Czapkiewicz \(2017\)](#) and [Hanauer and Lauterbach \(2019\)](#) also investigate firm characteristics belonging to categories such as profitability, investment, intangibles, and trading frictions. Our study includes characteristics from all these groups, but machine learning models can also take non-linearities and interactions into account next to linear relationships.

Finally, our paper also contributes to the understanding of the source of return predictability from machine learning forecasts. [Avramov et al. \(2022\)](#) show that return forecasts from deep learning models for the U.S. extract their profitability mainly from difficult-to-arbitrage stocks and during high limits-to-arbitrage market states. The authors also argue that the performance of machine learning forecasts further deteriorates when microcaps are excluded and when reasonable transaction costs are considered. Similarly, [Leung et al. \(2021\)](#) find that the economic gains of a gradient boosting machine model for developed market stocks tend to be more limited and critically dependent on the ability to take risk and implement trades efficiently. Furthermore, [Cakici et al. \(2022a\)](#) document that machine learning strategies work best for small stocks, as well as in countries with many listed firms and high idiosyncratic risk. In our paper, we follow [Hou et al. \(2020\)](#) and exclude microcaps from our analysis. While we also find that both linear and machine learning models show higher predictability for stocks associated with higher limits to arbitrage, we also document that this effect is less pronounced for machine learning models. Furthermore, we also provide evidence that a positive and significant outperformance and six-factor alpha can be achieved even when accounting for transaction costs, short-selling constraints, and limiting the investment universe to big stocks only.

The remainder of the paper is structured as follows: Section 2 describes the data sources, sample composition, and utilized firm-level characteristics. Section 3 outlines our methodology for predicting returns with machine learning algorithms, portfolio construction, and benchmark models. Section 4 presents evidence of the superiority of more complex machine learning models, while Section 5 strives to understand the source of this superiority better. We provide our conclusions in Section 6.

## 2. Data

### 2.1. Stock market data

Our sample comprises data from emerging stock markets as classified by Morgan Stanley Capital International (MSCI). The accounting data is from Refinitiv Worldscope, and the stock market data is from Refinitiv Datastream. The sample period starts in July 1990 and ends in December 2021. Countries are included in the sample only in those years in which they are part of the MSCI Emerging Markets Index.<sup>4</sup> Furthermore, countries are only part of the final sample in those months for which at least 10 stock-month observations are available after applying screens. The following 32 countries meet these criteria: Argentina, Brazil, Chile, China, Colombia, Czech Republic, Egypt, Greece, Hungary, India, Indonesia, Israel, Jordan, Korea, Kuwait, Malaysia, Mexico, Morocco, Pakistan, Peru, Philippines, Poland, Portugal, Qatar, Russia, Saudi Arabia, South Africa, Sri Lanka, Taiwan, Thailand, Turkey, and the UAE.<sup>5</sup>

We apply several static and dynamic screens to ensure that our sample comprises exclusively of common stocks and provides the highest data quality. First, we identify stocks using Refinitiv Datastream constituent lists, particularly Refinitiv Worldscope lists, research lists, and dead lists (to eliminate survivorship bias). Following [Ince and Porter \(2006\)](#), [Griffin et al. \(2010\)](#), [Schmidt et al. \(2019\)](#), [Hanauer \(2020\)](#), we eliminate non-common equity stocks through generic and country-specific static screens. Furthermore, we apply several dynamic screens to stock returns and prices to exclude erroneous and illiquid observations. [Appendix A](#) provides a detailed description of the constituent lists and the associated static and dynamic screens. Furthermore, we require stocks to have market capitalization data for the previous month.

We follow the size group methodology of [Fama and French \(2008\)](#), [Fama and French \(2012\)](#), [Fama and French \(2017\)](#) and [Hanauer and Lauterbach \(2019\)](#) and assign stocks into three size groups (micro, small, and big) for each country and month. Big stocks are the largest stocks, which together account for 90% of a country's aggregated market capitalization. Small stocks comprise the next 7% of aggregated market capitalization (so that big and small stocks together account for 97% of the aggregated market size of a country). Microcaps comprise the remaining 3%.<sup>6</sup> Although micro stocks represent only 3% of the total market capitalization of our emerging

<sup>4</sup> See <https://www.msci.com/market-classification> for details.

<sup>5</sup> The Chinese sample includes only non "A"-shares to proxy the investment universe for an international investor, as the China A-share market was only accessible to local investors for the majority of our sample period (cf., [Jansen et al., 2021](#)).

<sup>6</sup> To distinguish between these size groups, [Fama and French \(2008\)](#) use the 20th and 50th percentiles of end-of-June market cap on NYSE stocks as size breakpoints for the U.S. market, which on average are bigger than AMEX or NASDAQ stocks. However, these breakpoints are applied to all (NYSE, AMEX, and NASDAQ) stocks. For international markets, [Fama and French \(2012\)](#) and [Fama and French \(2017\)](#) propose to calculate breakpoints based on aggregated market capitalization, as we do.

market universe, they account for 67% of the number of stocks, which is similar to the proportion reported in Fama and French (2008) and Hanauer (2020) for the U.S. and developed markets, respectively. To prevent our results from being driven by microcaps, we follow Hou et al. (2020) and Hanauer and Lauterbach (2019) and exclude them. Finally, we cap the market capitalization of each stock within each month by its 99th percentile to avoid our results being driven by erroneous data and a few mega-caps.

We calculate returns from the total return index in USD. Following Jacobs (2016) and Hanauer and Lauterbach (2019), we win-sort all returns each month within a country at 0.1% and 99.9% to eliminate potential errors. To calculate the excess returns, we obtain the risk-free rate from Kenneth R. French's homepage.<sup>7</sup>

The result is a comprehensive dataset spanning 15,152 unique stocks and more than 1.42 million stock-month observations. Table 1 depicts the descriptive statistics for the final sample.

## 2.2. Firm-level characteristics

The 36 firm-level characteristics in this study are analogous to those in Kelly et al. (2019) and Windmüller (2022) and constructed using data from Refinitiv Datastream and Worldscope. Appendix B outlines the detailed construction of the characteristics. We follow Windmüller (2022) and substitute the daily bid-ask spreads with the daily version of Amihud (2002) illiquidity as a proxy for trading

**Table 1**

**Summary statistics by country.** This table presents summary statistics for the 32 countries of our sample. Column 1 reports the country names, and Columns 2, 3, 4, and 5 report the total, minimum, mean, and maximum number of firms per country. Columns 6 and 7 state the average mean and median size per country-month. Column 8 shows the average total size per country-month and column 9 reports these values in percentage of the respective total across countries. Size is measured as market capitalization in million USD. The last two columns report the actual beginning and ending dates during which each country is included in our sample.

Country	Total No. firms	Min No. firms	Mean No. firms	Max No. firms	Mean size	Median size	Total size	% of total size	Start date	End date
Argentina	96	11	30	45	832	376	25467	1.12	91-05	21-12
Brazil	289	17	65	154	3121	1500	247409	4.08	94-09	21-12
Chile	201	53	74	102	1740	826	125716	4.26	90-07	21-12
China	28	10	16	24	2772	1292	45063	0.10	16-05	21-12
Colombia	50	14	19	25	2791	2178	53710	0.95	94-07	21-12
Czech Rep.	89	10	36	77	662	234	13541	0.24	97-07	05-08
Egypt	199	51	81	123	577	213	46194	0.59	01-07	21-12
Greece	334	37	90	224	471	178	39442	1.56	90-07	21-12
Hungary	41	10	12	22	1921	486	22053	0.44	97-07	21-12
India	2238	356	593	893	1242	334	788478	13.09	94-07	21-12
Indonesia	649	35	150	296	1019	324	181003	4.50	90-07	21-12
Israel	634	173	245	331	270	60	62659	1.31	95-07	10-06
Jordan	161	10	98	119	328	70	32467	0.08	06-04	09-06
Korea	2972	394	803	1343	622	134	572232	12.87	92-07	21-12
Kuwait	81	73	75	78	1504	396	113350	0.02	21-07	21-12
Malaysia	1173	177	389	534	641	154	242228	9.97	90-07	21-12
Mexico	181	25	54	70	2692	1316	156068	4.92	90-07	21-12
Morocco	57	24	31	37	1323	658	42758	0.41	01-07	14-06
Pakistan	362	73	139	205	190	68	28089	0.50	94-07	21-12
Peru	103	17	28	40	1100	643	31521	0.61	94-07	21-12
Philippines	270	23	80	113	1087	433	101378	2.70	90-07	21-12
Poland	591	26	135	232	653	143	100667	1.74	95-07	21-12
Portugal	99	35	51	60	394	155	18182	0.84	90-07	98-06
Qatar	32	25	27	29	4969	2880	134437	0.42	14-07	21-12
Russia	232	10	54	102	4711	1958	285849	3.77	98-07	21-12
Saudi Arabia	89	34	52	84	8189	4936	398893	0.36	19-07	21-12
South Africa	517	80	131	240	2445	1101	274534	6.31	95-07	21-12
Sri Lanka	150	88	98	105	15	7	1530	0.03	94-07	01-06
Taiwan	1912	339	743	978	772	223	596406	11.36	97-07	21-12
Thailand	823	133	237	387	728	188	194393	6.48	90-07	21-12
Turkey	430	50	134	237	862	243	125222	3.77	90-07	21-12
UAE	69	33	43	49	4703	1653	201327	0.61	14-07	21-12
EM	15152	594	3763	5690	901	209	3909693	100.00	90-07	21-12

frictions. As shown by Fong et al. (2017), the Amihud (2002) illiquidity measure increases the number of observations in the cross-section and is the best daily cost-per-dollar-volume proxy for international data.

The 36 characteristics are: assets-to-market (A2ME), total assets (AT), sales-to-assets (ATO), book-to-market (BEME), market beta

<sup>7</sup> See [http://mba.tuck.dartmouth.edu/pages/faculty/ken.french/data\\_library.html](http://mba.tuck.dartmouth.edu/pages/faculty/ken.french/data_library.html).

(Beta), cash-and-short-term-investment-to-assets (C), capital turnover (CTO), capital intensity (D2A), leverage (Debt2P), ratio of change in property, plants, and equipment to change in total assets (DPI2A), earnings-to-price (E2P), fixed costs-to-sales (FC2Y), cash flow-to-book (FreeCF), idiosyncratic volatility (IdioVol), investment (INV), market capitalization (LME), turnover (LTurnover), net operating assets (NOA), operating accruals (OA), operating leverage (OL), price relative to its 52-week high (P2P52WH), price-to-cost margin (PCM), profit margin (PM), gross profitability (Prof), Tobin's Q (Q), momentum ( $r_{12-2}$ ), intermediate momentum ( $r_{12-7}$ ), short-term reversal ( $r_{2-1}$ ), long-term reversal ( $r_{36-13}$ ), return on net operating assets (RNA), return on assets (ROA), return on equity (ROE), sales-to-price (S2P), the ratio of sales and general administrative costs to sales (SGA2S), unexplained volume (SUV), and Amihud (2002) illiquidity (Illiqu).

Moreover, in a robustness check, we add the following four characteristics that have been shown to be strong return predictors for emerging markets (Hanauer and Lauterbach, 2019): monthly updated book-to-market ( $BEME_m$ , Asness and Frazzini, 2013), composite equity issuance (CEI, Daniel and Titman, 2006), cash flow-to-price (CF2P, Lakonishok et al., 1994), and gross profitability-to-assets (GP2A, Novy-Marx, 2013).

We do not exclude financial firms but set the following characteristics as missing as they are not meaningfully defined for financials: ATO, C, D2A, DPI2A, FC2Y, FreeCF, CF2P, GP2A, OA, PCM, PM, Prof, RNA, SGA2S, and NOA.

Following Freyberger et al. (2020), Gu et al. (2020), and Leippold et al. (2022), we rank all stock characteristics cross-sectionally for each month and country into the [-1,1] interval to limit the effect of outliers. These country-based ranks aim to address the impact of different accounting standards across countries, particularly in the earlier part of the sample period, and thus account for cross-country differences in characteristics. In case of missing characteristics, we replace them with a 0 to ensure broad cross-sectional coverage. Balance sheet data from the fiscal year ending in calendar year t-1 is used from end-of-June in year t to end-of-May in year t + 1 to predict stock returns from July in year t to end-of-June in year t + 1.

### 3. Methodology

#### 3.1. Return prediction using machine learning

Rasekhschaffe and Jones (2019) stress that domain knowledge is essential to structure the forecasting problem in a way that increases the signal-to-noise ratio. As we are interested in the cross-section of stock returns and rank stocks in portfolio sorts later in a country-neutral manner, we aim to forecast the outperformance of a stock relative to its country market return. Therefore, we define the abnormal return of a stock  $i, i = 1, \dots, N$  in month  $t, t = 1, \dots, T$  in the country  $c, c = 1, \dots, C$  as

$$r_{i,t,c}^{abn} = r_{i,t,c} - Mkt_{t,c}, \tag{1}$$

where  $r_{i,t,c}$  is the return of stock  $i$  in month  $t$  of country  $c$  and  $Mkt_{t,c}$  is the value-weighted market return in month  $t$  of country  $c$ .

Following Gu et al. (2020), we employ a general additive prediction model to describe the one-month-ahead abnormal return of a stock  $r_{i,t+1,c}^{abn}$ , which can be written as

$$r_{i,t+1,c}^{abn} = E_t[r_{i,t+1,c}^{abn} | x_{i,t}] + \epsilon_{i,t+1,c}, \tag{2}$$

where  $E_t[r_{i,t+1,c}^{abn} | x_{i,t}]$  is the conditional expected abnormal return of stock  $i$  in month  $t$  for month  $t + 1$  given a vector of stock-specific  $p$  characteristics known at month  $t, x_{i,t} \in \mathbb{R}^p$ , and  $\epsilon_{i,t+1,c}$  is the prediction error term. Our objective is to estimate the expected abnormal return by using an unknown function  $f^*, f^* : \mathbb{R}^p \rightarrow \mathbb{R}$ , which estimates the expected returns independently of any other information besides the vector of  $p$  stock-specific characteristics available in month  $t$ :

$$E_t[r_{i,t+1,c}^{abn} | x_{i,t}] = f^*(x_{i,t}). \tag{3}$$

In the case of supervised machine learning, the unknown function  $f^*(x)$  is approximated by some function  $f(x, \theta, \rho)$ , which is parameterized by a vector of coefficients  $\theta$  and a set of hyperparameters  $\rho$ . While  $\theta$  is directly derived from the underlying training data with respect to  $\rho$  and a specific loss function  $L, \rho$  itself depends on the user input but is optimized concerning  $L$  based on available data. The exact functional form of  $f$  depends on the family and can be either linear or non-linear, parametric or non-parametric.

For this paper, we build on Rasekhschaffe and Jones (2019), Gu et al. (2020), Tobek and Hronec (2020), Drobetz and Otto (2021), and Leippold et al. (2022) to select a representative amount of machine learning models from the finance literature. We analyze the predictive power of nine different algorithms: ordinary least squares (OLS) regression, elastic net (ENet), gradient-boosted regression trees (GBRT), random forest (RF), and neural networks with one to five layers ( $NN_1, NN_2, NN_3, NN_4, NN_5$ ). We also investigate the performance of an ensemble of the five different neural networks ( $NN_{1-5}$ ) and the average combination of the more advanced machine learning methods (ENS): GBRT, RF, and  $NN_{1-5}$ . We provide a more detailed description of the models in Appendix C.

Besides the model selection, we also follow the standard approach in the literature (Gu et al., 2020; Leippold et al., 2022) for selecting the hyperparameter range, the training of the models, and the performance evaluation. One of the most crucial things when estimating the different machine learning models is to avoid data leakage. This happens when information exceeding the training dataset is used to create the model. Therefore, we divide our data into three disjoint periods, which always maintain the temporal ordering: the training, validation, and testing samples. We first estimate the models for a range of hyperparameters based on the training data. Next, we determine the respective loss of each hyperparameter set and model in the validation sample. The optimal

hyperparameter set minimizes the individual model's respective loss function. Afterward, we retrain the model with the optimal hyperparameter set on the combined training and validation data. Next, the models are used to predict the monthly returns for the test dataset. We describe an example of this procedure for the first two years in our sample: we first estimate the models for a range of hyperparameters based on the training data from July 1990 to December 1995. Afterward, we determine the best hyperparameters through the validation sample from January 1996 to December 2001. Finally, the model is retrained with the optimal hyperparameter using the data from July 1990 to December 2001 and evaluated in the testing sample using data from January 2002 to December 2002. To test our models from January 2003 to December 2003, we extend the training sample by one year (July 1990 to December 1996) and roll the validation sample forward by one year (January 1997 to December 2002). This procedure ensures that no future information is leaked from a previous period. Since machine learning models are computationally intensive, we retrain them only once at the end of every year but do the prediction every month using the latest model and data. [Appendix C.5](#) summarizes the hyperparameter tuning schemes for each model.

### 3.2. Machine learning portfolios

We mainly rely on portfolio performance analysis to evaluate the predictive performance of the different machine learning models. For a given machine learning model, we follow the following approach: At the end of each month  $t$ , we predict the next month's abnormal return ( $\hat{r}_{i,t,c}^{abn}$ ), which we use for sorting stock into quintiles. To avoid that small stocks or certain countries dominate our results, we estimate the quintile breakpoints for each country separately based on big stocks as recommended in [Hou et al. \(2020\)](#) and applied in [Hanauer and Lauterbach \(2019\)](#). Furthermore, the machine-learning-based signals should not only contain information on the return predictability in equal-weighted sorts, which may be driven by smaller stocks, but also in value-weighted sorts, which are dominated by larger stocks. Finally, we construct a zero-net investment portfolio (long-short) that goes long in the highest quintile portfolio and short in the lowest quintile portfolio. We reassign and rebalance all portfolios at the end of each month.

### 3.3. Benchmark factor models

To benchmark the results of the different machine learning portfolio sorts, we consider the [Fama and French \(2018\)](#) six-factor model, i.e., the [Fama and French \(2015\)](#) five-factor model with a cash-based profitability factor and augmented with the [Carhart \(1997\)](#) momentum factor. The corresponding six factors are market (*RMRF*), size (*SMB*, small minus big), value (*HML*, high minus low), profitability (*RMW*, robust minus weak), investment (*CMA*, conservative minus aggressive), and momentum (*WML*, winners minus losers). These factors are based on the same stock sample as the machine learning portfolios, i.e., we also exclude microcaps. Furthermore, we use regional versions of the factors for studying emerging market regions. [Appendix D](#) provides a detailed description of how the factors are constructed.

## 4. Empirical results

This section presents evidence on the application of various machine learning models in emerging markets. We begin by analyzing the out-of-sample  $R_{OOS}^2$  of individual stock returns. Subsequently, we evaluate the importance of different characteristics, the sensitivity of the predicted returns to various characteristics, and the sensitivity to the interaction effects of different characteristics. Next, we employ portfolio sorts to assess the economic gains of using different machine learning models. Finally, we investigate the impact of various methodological changes and the robustness of our findings in emerging market subregions.

### 4.1. Prediction performance

[Table 2](#) presents the out-of-sample  $R_{OOS}^2$  for our set of machine learning models, which measures the predictive power on the individual stock level. In Panel B, we show the [Newey and West \(1987\)](#) adjusted [Diebold and Mariano \(1995\)](#) test statistics to compare the out-of-sample stock-level prediction performance between each machine learning model. We measure the pooled out-of-sample  $R_{OOS}^2$  in Panel A as:

$$R_{OOS}^2 = 1 - \frac{\sum_t \sum_i (r_{i,t,c}^{abn} - \hat{r}_{i,t,c}^{abn})^2}{\sum_t \sum_i (r_{i,t,c}^{abn})^2}. \quad (4)$$

The first row in Panel A of [Table 2](#) reports the  $R_{OOS}^2$  of the full sample. The *OLS* yields a benchmark  $R_{OOS}^2$  of 0.29%, which all other models improve except for the *ENet* ( $R_{OOS}^2$  of 0.18%). Since the *ENet* shrinks certain coefficients towards zero but does not consider interactions or non-linearities, it seems that this regularization does not increase the predictability. The *RF* and *GBRT* are superior to the *OLS*, producing fits of 0.40% and 0.52%, respectively. Only the  $NN_1$  underperforms the *GBRT* but outperforms all other linear and non-parametric models and yields a  $R_{OOS}^2$  of 0.49%. The  $NN_2$  to  $NN_5$  show  $R_{OOS}^2$  between 0.53% and 0.55%, with the  $NN_4$  performing the best. Creating an ensemble of neural networks ( $NN_{1-5}$ ) and an ensemble of the non-linear machine learning models (*ENS*) produces fits for both models of 0.60%.

**Table 2**

**Monthly out-of-sample stock-level prediction performance.** This table summarizes the monthly out-of-sample stock-level prediction performance using OLS (*OLS*), elastic net (*ENet*), random forest (*RF*), gradient boosted regression trees (*GBRT*), neural networks with 1 to 5 layers (*NN<sub>1</sub>–NN<sub>5</sub>*), an ensemble of the different neural networks (*NN<sub>1–5</sub>*), and an ensemble of the different non-linear machine learning algorithms (*ENS*). Panel A reports the monthly  $R^2_{OOS}$  statistics for the full sample and within subsamples that include only large stocks or small stocks. Panel B reports pairwise [Newey and West \(1987\)](#) adjusted Diebold-Mariano test statistics comparing the out-of-sample stock-level prediction performance among each machine learning model. Positive numbers indicate the column model outperforms the row model. Bold font indicates the difference is significant at 1% level or better for individual tests, and an asterisk indicates significance at the 1% level for 10-way comparisons via our conservative Bonferroni adjustment. The out-of-sample period is from January 2002 to December 2021.

	<i>OLS</i>	<i>ENet</i>	<i>RF</i>	<i>GBRT</i>	<i>NN<sub>1</sub></i>	<i>NN<sub>2</sub></i>	<i>NN<sub>3</sub></i>	<i>NN<sub>4</sub></i>	<i>NN<sub>5</sub></i>	<i>NN<sub>1–5</sub></i>	<i>ENS</i>
<b>Panel A: Percentage <math>R^2_{OOS}</math></b>											
Full Sample	0.29	0.18	0.40	0.52	0.49	0.53	0.53	0.55	0.54	0.60	0.60
Large firms	0.12	−0.01	0.25	0.30	0.19	0.24	0.27	0.31	0.31	0.34	0.38
Small firms	0.40	0.29	0.49	0.66	0.67	0.70	0.68	0.70	0.68	0.75	0.73
<b>Panel B: Between-model comparison of predictive performance</b>											
<i>OLS</i>		−2.13	<b>3.45*</b>	<b>6.35*</b>	<b>5.57*</b>	<b>5.42*</b>	<b>5.64*</b>	<b>6.05*</b>	<b>6.65*</b>	<b>7.34*</b>	<b>8.25*</b>
<i>ENet</i>			<b>4.53*</b>	<b>6.50*</b>	<b>5.91*</b>	<b>6.08*</b>	<b>6.29*</b>	<b>7.21*</b>	<b>7.01*</b>	<b>7.68*</b>	<b>8.19*</b>
<i>RF</i>				<b>6.80*</b>	<b>2.96</b>	<b>3.96*</b>	<b>4.38*</b>	<b>5.27*</b>	<b>4.57*</b>	<b>6.59*</b>	<b>12.65*</b>
<i>GBRT</i>					−0.72	0.71	0.68	1.74	1.00	<b>3.49*</b>	<b>7.59*</b>
<i>NN<sub>1</sub></i>						2.51	2.12	<b>3.24</b>	<b>2.85</b>	<b>9.19*</b>	<b>4.38*</b>
<i>NN<sub>2</sub></i>							−0.23	1.69	0.25	<b>6.01*</b>	2.37
<i>NN<sub>3</sub></i>								1.82	0.39	<b>5.86*</b>	<b>2.82</b>
<i>NN<sub>4</sub></i>									−1.42	<b>3.12</b>	1.45
<i>NN<sub>5</sub></i>										<b>5.30*</b>	<b>2.60</b>
<i>NN<sub>1–5</sub></i>											−0.45

A closer look at the second and third rows in Panel A of [Table 2](#) reveals an interesting pattern: in all the cases, the predictive performance is better for small firms than for large firms. The ensemble of neural networks (*NN<sub>1–5</sub>*) and the ensemble of non-linear machine learning models (*ENS*) yield a  $R^2_{OOS}$  of 0.34% and 0.38% for large firms and 0.75% and 0.73% for small firms, respectively.

Whereas Panel A measures the individual predictive performance of the different machine learning models, Panel B assesses the statistical significance of differences among the models using the [Newey and West \(1987\)](#) adjusted [Diebold and Mariano \(1995\)](#) test statistics ( $DM_{kj}$ ) comparing a column model (*k*) versus a row model (*j*). We compute the Newey-West adjusted Diebold-Mariano test statistics as:

$$\begin{aligned}
 MSFE_t^m &= \frac{1}{N_t} \sum_{i=1}^{N_t} (r_{i,t,c}^{abn} - \hat{r}_{i,t,c,m}^{abn})^2 \\
 d_{kj,t} &= MSFE_t^k - MSFE_t^j \\
 \bar{d}_{kj} &= \frac{1}{T} \sum_t d_{kj,t} \\
 DM_{kj} &= \frac{\bar{d}_{kj}}{\hat{\sigma}_{d_{kj},NW(4)}}
 \end{aligned}
 \tag{5}$$

where  $\hat{\sigma}_{d_{kj},NW(4)}$  is the [Newey and West \(1987\)](#) standard error of  $d_{kj,t}$  with four lags. The Diebold-Mariano test statistic is normally distributed with a mean of 0 and a standard deviation of 1 ( $\mathcal{N}(0, 1)$ ) with the null hypothesis that there exists no difference between the models, which allows us to map the magnitudes of the test statistic to *p*-values. Bold numbers indicate a significant difference between the models at the 1% level ( $DM \geq 2.60$ ). An asterisk indicates statistical significance at the 1% level for 10-way comparisons via the conservative Bonferroni adjustment, which increases the critical value to 3.33.

Except for the *ENet*, all machine learning models outperform the *OLS* and exceed the Bonferroni adjusted critical value of 3.33. The comparison between the *RF* and all other non-linear models yields a similar result, with the *GBRT* and all other neural networks, except *NN<sub>1</sub>*, outperforming the *RF*. For the *GBRT*, only the two machine-learning ensembles significantly improve prediction performance, as evidenced by a *DM* statistic of 3.49 and 7.59, respectively. The different neural networks with one to five layers do not differ much in their prediction performance. In the case of the *NN<sub>1</sub>*, the neural networks with four and five layers are superior. The two best-performing machine learning models are the two ensembles. While the ensemble of neural networks (*NN<sub>1–5</sub>*) significantly outperforms all other machine learning models, the ensemble of the trees and neural networks (*ENS*) exhibits statistically significant outperformance when compared to the *OLS*, *ENet*, *RF*, *GBRT*, *NN<sub>1</sub>*, *NN<sub>3</sub>*, and *NN<sub>5</sub>*.

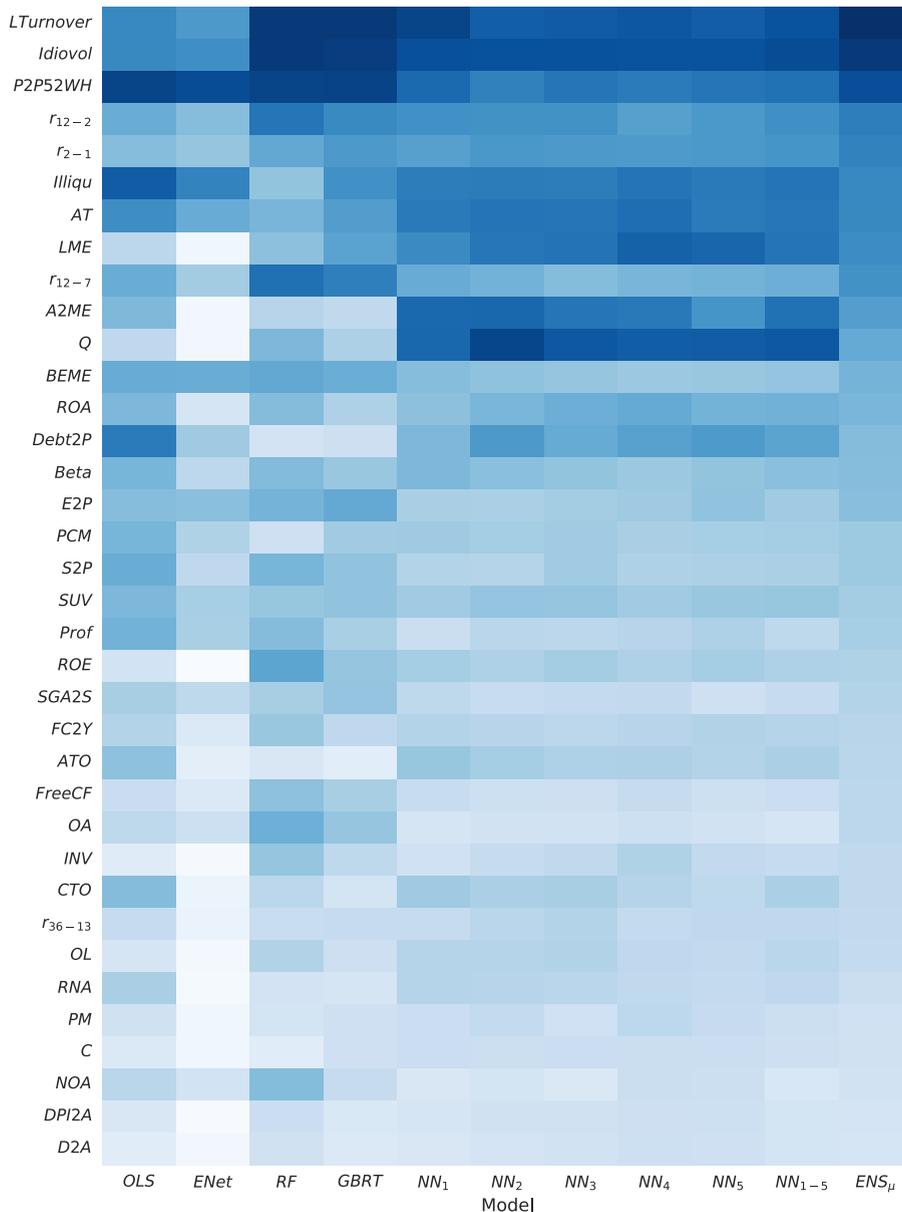
#### 4.2. Characteristics importance and marginal relationships

Next, we examine the importance of individual characteristics in predicting abnormal returns and the model-implied marginal

impact of individual characteristics on expected abnormal returns.

We determine the importance of each characteristic for each model by measuring the average reduction in  $R^2_{005}$  by setting each value of the particular characteristic to zero and keeping the remaining model estimates fixed. Fig. 1 visualizes the sum over the cross-sectional ranked characteristics for the different machine learning models.<sup>8</sup> A darker color in the figure indicates higher importance of the characteristic for the individual model, while a lighter color indicates lower importance for the  $R^2_{005}$ .

The most influential characteristics are similar among the different machine learning models, with turnover (*LTurnover*), idiosyncratic volatility (*Idiovol*), price relative to its 52-week high (*P2P52WH*, *Amihud (2002)* illiquidity (*Iliqu*), total assets (*AT*), market capitalization (*LME*), and market beta (*Beta*) from the trading frictions category; momentum ( $r_{12-2}$ ), short-term reversal ( $r_{2-1}$ ), and intermediate momentum ( $r_{12-7}$ ) from the past returns category; and assets-to-market (*A2ME*), Tobin's Q (*Q*), book-to-market (*BEME*),



**Fig. 1. Characteristic importance by model.** This figure shows the ranked characteristic importance for the variables in each model. Characteristic importance is an average over all training samples and importance within each model is normalized to sum to one.

<sup>8</sup> In addition, we present the most influential characteristics per model and the corresponding normalized importance in Fig. E1 in the appendix.

and leverage (*Debt2P*) from the value category all being among the top 15 characteristics. However, characteristics from the profitability and intangibles categories, except for return on asset (*ROA*), are not present among the top 15.

Fig. 2 visualizes the marginal impact of individual characteristics on expected abnormal returns for the *OLS*, *ENet*, *RF*, *GBRT*, and *NN<sub>1-5</sub>*. We predict the returns for each model and characteristic by iterating over the (-1,1) interval and holding all other characteristics fixed at zero. We do this for each time period and model individually and calculate the average predicted return among the different machine learning models. We select short-term reversal (*r<sub>2-1</sub>*), idiosyncratic volatility (*Idiovol*), turnover (*LTurnover*), and operating leverage (*OL*) as examples to visualize how the different machine learning models associate the underlying characteristic with the expected abnormal returns.

Inspecting the relationships in Fig. 2, we observe that all methods identify the well-known negative relationship between expected returns with short-term reversal (*r<sub>2-1</sub>*, top-left) or idiosyncratic volatility (*Idiovol*, top-right). While the two linear models are, per definition, restricted to linear relationships, we see that tree-based methods and neural networks identify more pronounced short-term reversal patterns in the extremes.<sup>9</sup> Similarly, these methods detect a relatively flat relationship for low and medium levels of idiosyncratic volatility (*Idiovol*) but an increasingly negative relationship for high idiosyncratic volatility, echoing the empirical results in Ang et al. (2006). The differences are even more pronounced for turnover (*LTurnover*, bottom-left). While both *OLS* and *ENet* find a positive slope, the two tree-based models, *RF* and *GBRT*, and the neural network ensemble, *NN<sub>1-5</sub>*, identify an inverted U-shape pattern: extreme positive and negative values of *LTurnover* are associated with lower expected return than the middle region in the interval, echoing the pattern documented in Freyberger et al. (2020). Such differences in marginal relationships can partly explain the divergence in the performance of linear and non-linear methods. However, we also observe that all methods agree on a nearly zero relationship between operating leverage (*OL*, bottom-right) and expected returns.

A significant advantage of the tree-based models and the different neural networks is that they can model complex interactions among the various characteristics. In Fig. 3, we illustrate how the *NN<sub>1-5</sub>* can model complex interactions between characteristics. Specifically, we show the sensitivity of the expected returns to pairwise interaction effects for Amihud (2002) illiquidity (*Illiqu*) and idiosyncratic volatility (*Idiovol*) with short-term reversal (*r<sub>2-1</sub>*) and market capitalization (*LME*) by varying both pairs of characteristics while holding the other predictors fixed. We choose *Illiqu* and *Idiovol* as they are prominent hard-to-value proxies (cf., Kumar, 2009) and *r<sub>2-1</sub>* and *LME* as they are two main control characteristics in the asset pricing literature.

The upper-left figure illustrates that the difference between high and low previous month returns is the most substantial for very

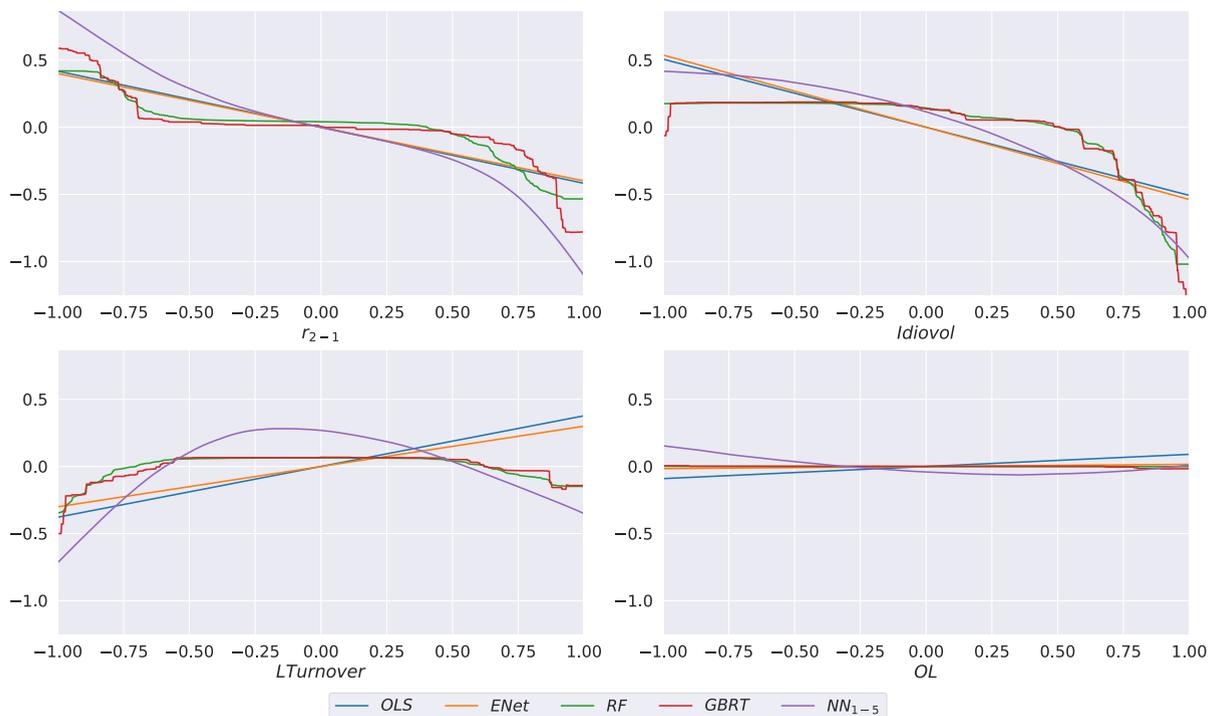
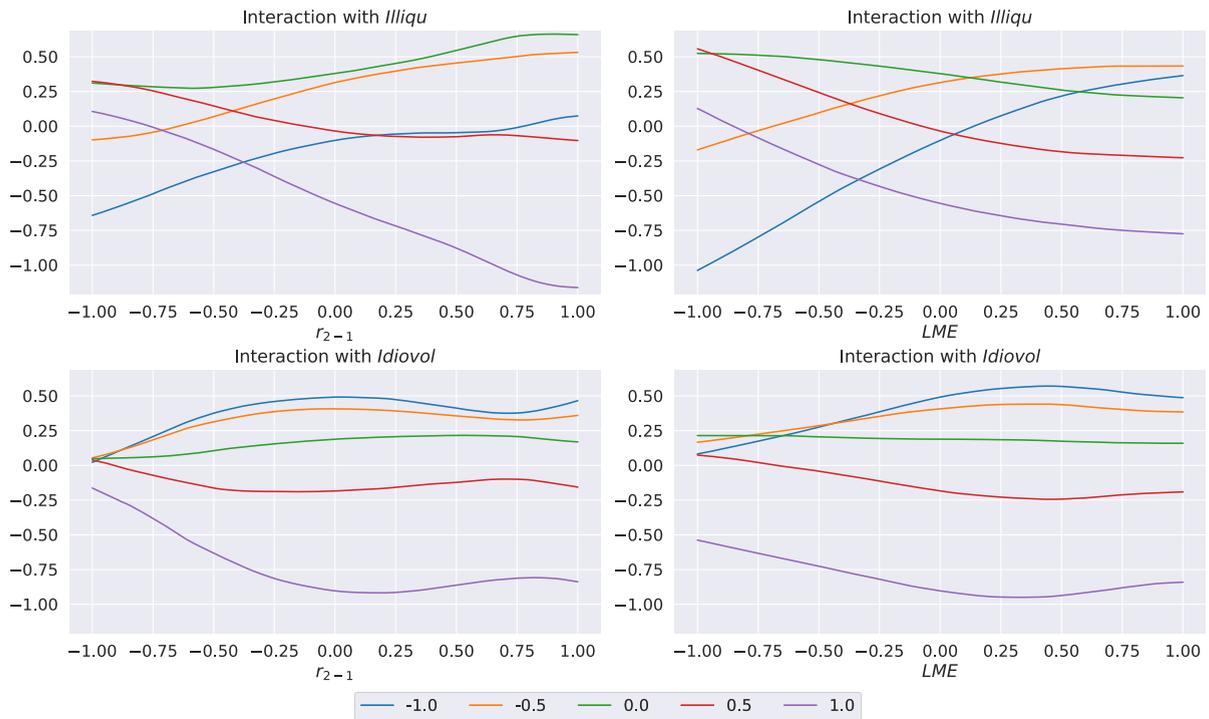


Fig. 2. Marginal association between expected returns and characteristics. The figure shows the sensitivity of expected returns (vertical axis) to the four following individual characteristics (holding all other covariates fixed at their median values): short-term reversal (*r<sub>2-1</sub>*, top-left), idiosyncratic volatility (*Idiovol*, top-right), turnover (*LTurnover*, bottom-left), and operating leverage (*OL*, bottom-right).

<sup>9</sup> This finding is consistent with the empirical pattern for short-term reversal deciles that can be found on Kenneth R. French's homepage.



**Fig. 3. Expected returns and characteristic interactions ( $NN_{1-5}$ ).** The figure shows the sensitivity of the expected returns (vertical axis) to interactions effects for four selected combinations in model  $NN_{1-5}$  (holding all other characteristics fixed at their median values of 0): Amihud (2002) illiquidity (*Illiqu*) and short-term reversal ( $r_{2-1}$ ) (top-left), Amihud (2002) illiquidity and market capitalization (*LME*) (top-right), idiosyncratic volatility (*Idiovol*) and short-term reversal (bottom-left), and idiosyncratic volatility and market capitalization (bottom-right).

illiquid stocks (purple line). In contrast, the lines remain mostly parallel for other values of *Illiqu*. This finding is consistent with the empirical observation for the interaction between short-term reversal with turnover reported in Medhat and Schmelming (2022). The upper-right figure depicts the interactions between a stock’s market capitalization and the Amihud (2002) illiquidity measure. For liquid firms (blue and orange line), the expected return increases with market capitalization. In contrast, the relationship is reversed for illiquid firms (red and purple line), implying that expected returns decrease for larger firms. The bottom-left figure reveals that the short-term reversal effect is most pronounced and S-shaped for risky stocks (purple line). In contrast, the reversal effect is concave for less risky stocks (blue and orange line), yielding significantly lower returns when the prior month’s returns are high. Finally, the bottom-right figure indicates that no strong interaction effects exist between *Idiovol* and *LME*.

4.3. Portfolio performance

Following our analysis of the predictive ability of the different machine learning methods for individual stock returns, we will now proceed with a general overview of the profitability of machine learning signal-based portfolios.

Table 3 displays the results of our analysis on equal- and value-weighted country-neutral quintile portfolio sorts using big-stock breakpoints. In Panel A and Panel D, we report the predicted monthly returns for the long-short quintile (Pred), the average monthly return for the long-short quintile (Avg), Newey and West (1987) adjusted *t*-statistics with four lags (*t*-stat), monthly standard deviations (SD), and Sharpe ratios (SR). Panel B and Panel E show the alphas ( $\alpha$ ), corresponding Newey and West (1987) adjusted *t*-statistics with four lags ( $t\text{-stat}_\alpha$ ), and  $R^2$  with respect to the Fama and French (2018) six-factor model:

$$r_{t,ML} = \alpha + \beta_1 RMRF_t + \beta_2 SMB_t + \beta_3 HML_t + \beta_4 RMW_t + \beta_5 CMA_t + \beta_6 WML_t + \epsilon_t, \tag{6}$$

where  $r_{t,ML}$  is the long-short quintile return in month *t*.

We additionally provide detailed results on every quintile in Table F1 in the appendix.<sup>10</sup>

Panel C and Panel F describe the maximum drawdowns (Max DD), the most negative monthly return (Max 1 M Loss), and the

<sup>10</sup> We also include the performance of a strategy that uses the equal-weighted (1/N) average of all standardized characteristics ( $\mu_{sign(c)}$ ) in this table. Thereby, characteristics are sorted in such a way that higher values correspond to higher expected returns. The performance of this simple linear combination is slightly worse (similar) than that of the other two linear strategies for equal-weighted (value-weighted) portfolios.

**Table 3**

**Drawdowns, turnover, and risk-adjusted performance of machine learning portfolios.** This table reports the out-of-sample performance of the different machine learning long-short portfolios. Stocks are sorted into country-neutral quintile portfolios based on their predicted returns for the next month. The sorting breakpoints are based on big stocks only, which are in the top 90% of a country's aggregated market capitalization. Panel A (Panel D) summarizes the quintile sort results from equal-weighting (value-weighting) and provides the predicted monthly returns for the long-short quintile (Pred), the average monthly returns of the long-short quintile (Avg), Newey and West (1987) adjusted t-statistics with 4 lags ( $t\text{-stat}_a$ ), their standard deviations (SD), and annualized Sharpe ratios (SR), respectively. Panel B (Panel E) reports the average Fama and French (2018) six-factor model alphas ( $\alpha_{FF6}$ ), corresponding Newey and West (1987) adjusted t-statistics with 4 lags ( $t\text{-stat}_a$ ), and corresponding  $R^2$  using equal-weighting (value-weighting). Panel C (Panel F) describes the maximum drawdowns (Max DD), the most negative monthly return (Max 1 M Loss), and the average monthly turnover in % of the equal-weighted (value-weighted) long-short portfolio. The sample period is from January 2002 to December 2021.

	OLS	ENet	RF	GBRT	NN <sub>1</sub>	NN <sub>2</sub>	NN <sub>3</sub>	NN <sub>4</sub>	NN <sub>5</sub>	NN <sub>1-5</sub>	ENS
<b>Panel A: Quintile sorts performance - Equal-weighted</b>											
Pred	1.93	1.97	1.50	1.80	2.61	2.60	2.41	2.29	2.25	2.30	1.71
Avg	1.38	1.20	1.60	1.82	1.89	1.91	1.84	1.86	1.85	1.88	1.86
$t\text{-stat}$	7.82	6.83	9.33	11.57	14.01	15.75	14.81	13.82	13.50	13.50	11.79
SD	2.04	2.12	2.04	1.88	1.68	1.58	1.60	1.66	1.69	1.72	1.87
SR	2.34	1.96	2.71	3.35	3.91	4.21	4.00	3.89	3.78	3.79	3.44
<b>Panel B: Risk-adjusted performance - Equal-weighted</b>											
$\alpha_{FF6}$	0.97	0.83	1.19	1.40	1.47	1.55	1.49	1.48	1.44	1.46	1.43
$t\text{-stat}_a$	8.02	6.94	14.10	15.65	15.67	19.02	16.93	16.72	15.79	15.81	15.66
$R^2$	62.42	55.79	59.81	60.89	53.21	48.65	52.70	54.66	56.15	55.67	58.17
<b>Panel C: Drawdowns and turnover - Equal-weighted</b>											
Max DD (%)	26.35	26.23	21.69	18.84	16.70	13.45	16.00	16.07	17.61	17.82	19.04
Max 1 M loss (%)	13.97	12.96	10.53	10.70	9.37	7.65	10.20	10.17	10.25	10.75	10.68
Turnover (%)	89.27	96.38	89.61	97.39	101.87	102.02	100.80	99.21	99.50	99.72	95.77
<b>Panel D: Quintile sorts performance - Value-weighted</b>											
Pred	1.85	1.89	1.39	1.61	2.30	2.21	2.04	1.94	1.93	1.97	1.52
Avg	0.84	0.73	0.99	1.06	1.04	1.12	1.12	1.20	1.17	1.21	1.21
$t\text{-stat}$	4.64	4.01	5.28	6.14	7.00	9.47	7.91	8.35	8.17	8.55	7.04
SD	2.22	2.36	2.32	2.17	1.95	1.75	2.01	1.97	1.87	1.98	2.20
SR	1.31	1.07	1.48	1.69	1.85	2.23	1.93	2.11	2.17	2.12	1.91
<b>Panel E: Risk-adjusted performance - Value-weighted</b>											
$\alpha_{FF6}$	0.28	0.27	0.47	0.57	0.57	0.71	0.66	0.73	0.71	0.72	0.67
$t\text{-stat}_a$	2.72	2.28	5.24	6.73	4.83	9.16	6.55	7.76	8.21	8.26	8.29
$R^2$	68.25	56.39	67.61	68.50	52.97	48.50	51.79	58.50	59.39	56.86	67.17
<b>Panel F: Drawdowns and turnover - Value-weighted</b>											
Max DD (%)	30.49	31.45	31.07	26.54	23.63	15.44	23.19	21.81	20.81	20.36	25.28
Max 1 M loss (%)	16.60	17.82	14.46	14.73	16.45	9.58	17.36	15.90	12.83	14.81	14.81
Turnover (%)	91.28	97.18	90.46	101.10	103.81	106.35	106.02	104.35	104.43	101.46	96.85

average monthly percentage change in holdings (TO) of different machine learning-based long-short portfolios. We define maximum drawdowns as

$$\text{Max DD} = \max_{0 \leq t_1 \leq t_2 \leq T} (Y_{t_1} - Y_{t_2}), \tag{7}$$

where  $Y_t$  is the cumulative log return from date 0 through  $t$ . The strategy's average monthly turnover is defined as

$$\text{TO} = \frac{1}{T} \sum_{t=1}^T \left( \sum_{i=1}^{N_t} w_{i,t+1} - \frac{w_{i,t}(1 + r_{i,t+1})}{1 + \sum_{j=1}^{N_t} w_{j,t} r_{j,t+1}} \right), \tag{8}$$

where  $w_{i,t}$  is the weight of stock  $i$  in the portfolio at time  $t$ .

We start by analyzing the equal-weighted long-minus-short quintile returns in Panel A of Table 3. All machine learning models yield positive and highly significant long-short returns. The order is similar to the monthly out-of-sample stock-level prediction performance in Table 2. The linear methods OLS and ENet yield a monthly return of 1.38% ( $t\text{-stat}$  7.82) and 1.20% ( $t\text{-stat}$  6.83), respectively. However, the tree-based methods RF and GBRT exhibit even higher long-short returns of 1.60% ( $t\text{-stat}$  9.33) and 1.82% ( $t\text{-stat}$  11.57), which themselves are outperformed by the neural networks with returns between 1.84% ( $NN_3$ ) and 1.91% ( $NN_2$ ) and  $t$ -statistics between 13.50 ( $NN_5$ ) and 15.75 ( $NN_2$ ). The ensemble of the different neural networks ( $NN_{1-5}$ ) yields a similar performance as  $NN_5$ , and the ensemble of the tree-based methods and neural networks has a performance similar to the GBRT.

The risk-adjusted performance displayed in Panel B leads to the same order as the raw long-short returns. However, the increase in

the six-factor alpha for the machine learning models compared to the linear models is even more pronounced as the six-factor model has less explanatory power. Furthermore, Panel C reveals that the neural network portfolios exhibit a smaller maximum drawdown and maximum one-month loss than the linear and tree-based models. The maximum drawdown (worst one-month return) in the case of the ensemble of neural networks is 17.82% (10.75%), whereas this number is 26.35% (13.97%) for the *OLS*. The superior performance of the machine learning models comes at the cost of a somewhat higher turnover. However, compared to the performance gains, this turnover increase from 89.27% for *OLS* to values between 89.61% for *RF* and 102.02 for *NN<sub>2</sub>* is relatively small.

Turning to the results for value-weighted portfolios in Panels D to E of Table 3 reveals identical qualitative conclusions, but the return spreads, *t*-statistics, and Sharpe ratios are substantially lower. Although the return forecasts derived from linear models already lead to economically and statistically significant long-short mean returns and six-factor alphas, the tree-based methods and neural networks do even better. Again, the neural network with two layers exhibits the highest *t*-statistics and Sharpe ratios while suffering from the mildest drawdowns. Comparing the ensemble of machine learning methods (*ENS*) with the linear *OLS* regressions shows performance gains of roughly 50% for the raw quintile returns and even higher for the risk-adjusted performance. In sum, allowing for non-linearities and interactions also leads to economically superior out-of-sample returns compared to traditional linear models, as summarized in Fig. 4.

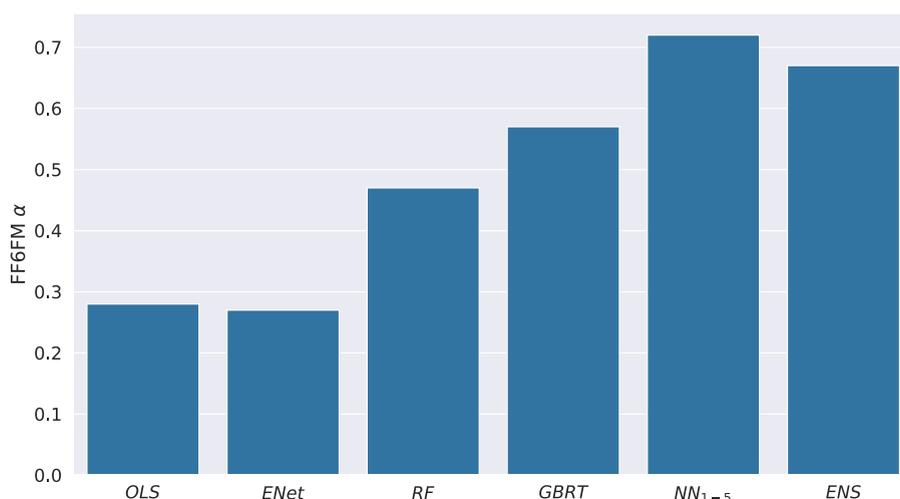
Fig. 5 illustrates the results of Table 2 by plotting the equal-weighted and value-weighted cumulative performance of selected long-short strategies. We additionally include the cumulative performance for the long and short sides for select strategies in Appendix E2. Notably, the performance of our strategies does not predominantly stem from the short side, which would raise investability concerns due to shorting frictions.

Using a value-weighted portfolio strategy, *RF* initially dominates the other methods, while the outperformance of *GBRT* and *NN<sub>1-5</sub>* mainly stems from the period after 2009. As the *ENS* comprises all three methods, we observe a rather consistent outperformance versus *OLS* that is not driven by a particular period. In the case of equal-weighted portfolios, there are only small differences between the portfolio returns of *GBRT*, *NN<sub>1-5</sub>*, and *ENS* till 2021. As for the value-weighted portfolios, the machine learning methods outperform the linear approaches consistently over time. The model with the lowest cumulative return is the *ENet*, whereby the underperformance versus the *OLS* is mainly driven by the first years of the sample period. Besides a sharp drawdown in 2009, there are no other notable downturns for all approaches. The drawdown in 2009 probably stems from the models' exposure to momentum that exhibited a momentum crash at that time (Daniel and Moskowitz, 2016; Hanauer and Windmüller, 2023). The recent global shock due to the COVID-19 pandemic in early 2020 did not lead to a significant portfolio-level downturn.

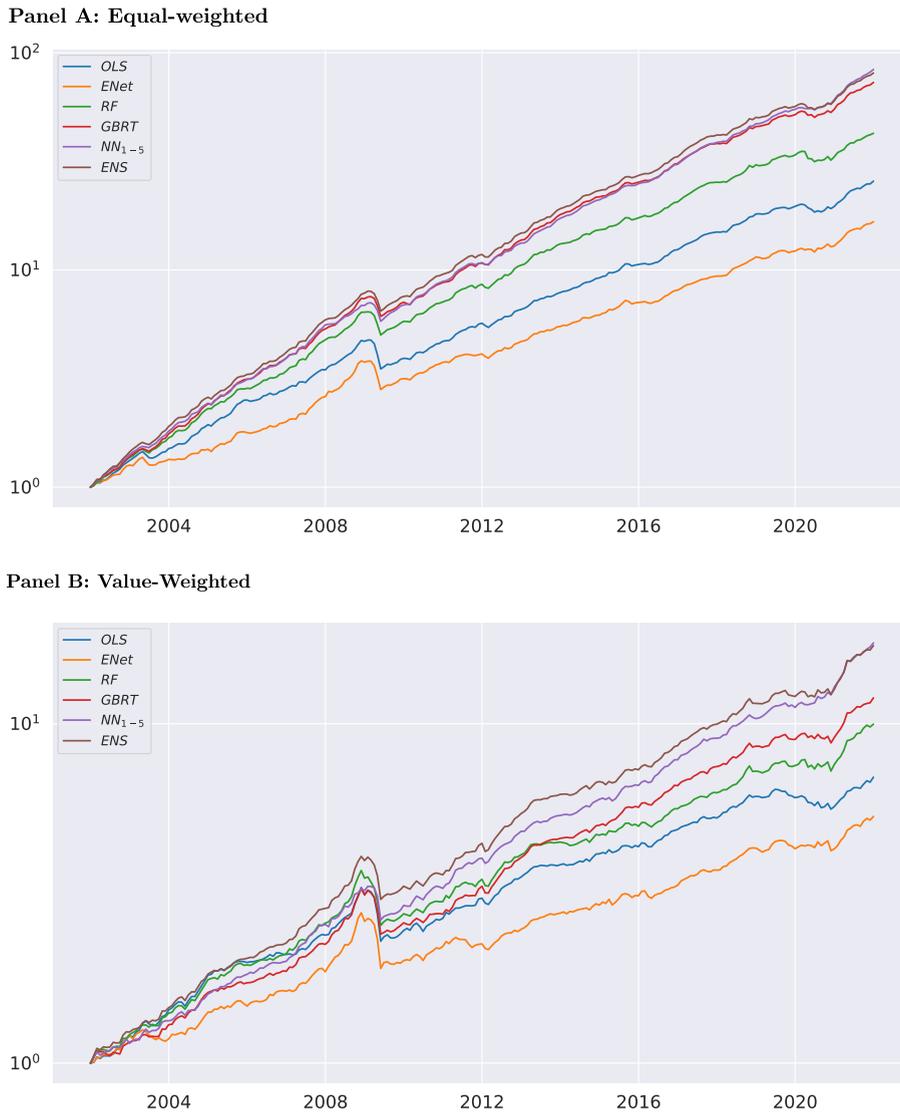
#### 4.4. Robustness

To check the robustness of the results presented above, we will investigate (i) the impact of various methodological changes and (ii) the robustness within emerging market subregions.

Table 4 summarizes the robustness tests for methodological changes. We include various performance indicators for our equal-weighted and value-weighted machine learning portfolio strategies. However, we will only compare the performance of the benchmark *OLS* model to the *ENS* model. We select the *ENS* to be not driven by a look-ahead bias regarding the model selection and its portfolio performance. Besides the individual long-short return and the six-factor model of the two machine learning models, we include the results of the following two regressions in the last two rows of each panel:



**Fig. 4. Fama and French (2018) six-factor model alphas.** This figure shows the Fama and French (2018) six-factor models alphas for various machine learning long-short portfolios. Stocks are sorted into country-neutral and value-weighted quintiles based on their predicted returns for the next month. The sorting breakpoints are based on big stocks only, which are in the top 90% of a country's aggregated market capitalization. The sample period is from January 2002 to December 2021.



**Fig. 5. Cumulative return of machine learning portfolios.** The figure shows the cumulative log returns of long-short quintile portfolios sorted on the out-of-sample machine learning return forecasts. Panel A shows equal-weighted returns, while Panel B shows value-weighted returns. The sample period is from January 2002 to December 2021.

$$\begin{aligned}
 r_{LS,t,ENS} &= \alpha + \beta_{OLS} r_{LS,t,OLS} + \epsilon_t, \\
 r_{LS,t,OLS} &= \alpha + \beta_{ENS} r_{LS,t,ENS} + \epsilon_t.
 \end{aligned}
 \tag{9}$$

A positive and significant alpha indicates that the returns of the strategy on the right-hand side cannot fully explain the portfolio returns on the left-hand side.

The first two rows in Panel A show again our baseline result for *OLS* and *ENS*, as previously shown in Table 3. In addition, the last two rows of Panel A demonstrate that the *ENS* long-short portfolio spans the *OLS* long-short portfolio for both equal- and value-weighted portfolios, but the *OLS* portfolios cannot span *ENS* portfolios.

In Panel B, we construct our long-short trading strategy using decile instead of quintile sorts. By focusing on more extreme predicted abnormal returns and due to the monotonic increase among the portfolios, the equal-weighted and value-weighted long-short returns of the *OLS* increase to 1.84% (*t*-stat 10.09) and 1.18% (*t*-stat 5.91), whereas the Fama and French (2018) six-factor alpha increases to 1.41% (*t*-stat 11.30) and 0.55% (*t*-stat 4.66). The *ENS* shows an increase in the return to 2.50% (*t*-stat 13.93) and 1.66% (*t*-stat 8.12) as well as in the risk-adjusted return to 2.02% (*t*-stat 18.31) and 1.10% (*t*-stat 10.30). Therefore, both *OLS* and *ENS* show stronger results when using decile sorts. Still, the increase in returns of the *ENS* is higher than the *OLS*, resulting in a larger  $\alpha$  when regressing the *ENS* on the *OLS* compared to Panel A.

The feature set for the robustness test presented in Panel C includes the additional predictive characteristics described in Hanauer

**Table 4**

**Robustness.** This table reports robustness tests for the out-of-sample performance of equal- and value-weighted long-short portfolios. All stocks are sorted into country-neutral quintile portfolios based on their predicted returns for the next month. We investigate predictions from a linear OLS model and an ensemble (*ENS*) of non-linear machine learning models (*RF*, *GBRT*, and  $NN_{1-5}$ ). The sorting breakpoints are based on big stocks only, which are in the top 90% of a country's aggregated market capitalization. Panel A summarizes the baseline results as presented in Table 3. Panel B reports results on using decile sorts. Panel C uses an extended feature set following Hanauer and Lauterbach (2019). Panel D applies a feature selection before training the machine learning algorithms. Panel E uses predictions stemming from machine learning algorithms only trained on developed market data. Panel F excludes the high-turnover characteristics  $IdioVol$ ,  $LTurnover$ ,  $r_{2-1}$ ,  $SUV$ ,  $Illiqu$  from the feature set. Panel G shows the results for models trained on emerging market subregions. The first two rows of each panel provide the average monthly returns of the long-short quintile (*Avg*), corresponding *t*-statistics (*t*), the average Fama and French (2018) six-factor alphas ( $\alpha$ ), corresponding *t*-statistics ( $t_\alpha$ ), and  $R^2$ . The next two rows show spanning alphas ( $\alpha$ ), corresponding *t*-statistics ( $t_\alpha$ ), and  $R^2$  when regressing the long-short *ENS* returns on *OLS* returns and vice versa. All *t*-statistics are calculated using Newey and West (1987) adjusted standard errors with 4 lags. The sample period is from January 2002 to December 2021.

	Equal-weighted					Value-weighted				
	Avg	<i>t</i>	$\alpha$	$t_\alpha$	$R^2$	Avg	<i>t</i>	$\alpha$	$t_\alpha$	$R^2$
<b>Panel A: Baseline</b>										
<i>OLS</i>	1.38	7.82	0.97	8.02	62.42	0.84	4.64	0.28	2.72	68.25
<i>ENS</i>	1.86	11.79	1.43	15.66	58.17	1.21	7.04	0.67	8.29	67.17
<i>ENS</i> ~ <i>OLS</i>			0.73	9.01	80.28			0.49	7.83	74.66
<i>OLS</i> ~ <i>ENS</i>			-0.44	-2.10	80.28			-0.22	-1.50	74.66
<b>Panel B: Decile sorts</b>										
<i>OLS</i>	1.84	10.09	1.41	11.30	52.39	1.18	5.91	0.55	4.66	62.37
<i>ENS</i>	2.50	13.93	2.02	18.31	54.16	1.66	8.12	1.10	10.30	57.37
<i>ENS</i> ~ <i>OLS</i>			1.01	7.10	71.88			0.78	6.00	56.06
<i>OLS</i> ~ <i>ENS</i>			-0.38	-2.76	71.88			-0.07	-0.36	56.06
<b>Panel C: Extended feature set</b>										
<i>OLS</i>	1.51	9.51	1.14	11.05	52.93	0.87	5.22	0.36	3.17	55.86
<i>ENS</i>	1.96	13.14	1.57	19.57	56.22	1.22	6.80	0.71	7.55	63.06
<i>ENS</i> ~ <i>OLS</i>			0.69	9.72	80.98			0.45	5.37	73.16
<i>OLS</i> ~ <i>ENS</i>			-0.38	-2.43	80.98			-0.13	-1.26	73.16
<b>Panel D: Feature selection</b>										
<i>OLS</i>	1.36	8.03	0.97	8.94	61.23	0.82	4.51	0.28	2.61	65.34
<i>ENS</i>	1.83	12.31	1.43	17.52	58.83	1.23	7.36	0.73	8.71	63.59
<i>ENS</i> ~ <i>OLS</i>			0.75	8.96	80.53			0.60	9.09	69.83
<i>OLS</i> ~ <i>ENS</i>			-0.50	-2.34	80.53			-0.29	-1.60	69.83
<b>Panel E: Trained on developed markets</b>										
<i>OLS</i>	1.29	6.71	0.93	6.76	62.40	0.89	4.67	0.38	3.37	68.17
<i>ENS</i>	1.67	10.55	1.23	10.12	59.97	1.20	6.64	0.62	5.31	61.15
<i>ENS</i> ~ <i>OLS</i>			0.72	6.99	79.14			0.43	4.75	74.18
<i>OLS</i> ~ <i>ENS</i>			-0.50	-3.84	79.14			-0.15	-1.20	74.18
<b>Panel F: Excluding short-term feature set</b>										
<i>OLS</i>	1.36	7.12	0.89	8.85	63.66	0.77	4.25	0.24	3.37	74.39
<i>ENS</i>	1.59	9.13	1.17	11.62	59.27	1.00	5.56	0.46	5.91	70.92
<i>ENS</i> ~ <i>OLS</i>			0.45	7.31	88.62			0.32	3.93	82.31
<i>OLS</i> ~ <i>ENS</i>			-0.32	-4.99	88.62			-0.16	-2.18	82.31
<b>Panel G: Subregional training</b>										
<i>OLS</i>	1.09	6.29	0.77	6.23	53.92	0.78	4.42	0.29	2.57	56.92
<i>ENS</i>	1.35	8.88	0.97	9.75	57.22	0.97	5.95	0.44	4.59	58.10
<i>ENS</i> ~ <i>OLS</i>			0.49	7.78	77.56			0.28	4.46	75.77
<i>OLS</i> ~ <i>ENS</i>			-0.23	-1.76	77.56			-0.06	-0.51	75.77

and Lauterbach (2019). The *OLS* particularly benefits from this extended feature set. The average equal-weighted and value-weighted long-short return increase by 9% and 4%, while only the equal-weighted return of *ENS* increases by 5%. In the case of the value-weighted risk-adjusted return, the *OLS* alpha increases by 28% and the *ENS* alpha increases by 6%.

Reducing the number of characteristics by applying a lasso regression, i.e., feature selection, before training the machine learning models reduces the equal-weighted and value-weighted long-short returns as well as the equal-weighted risk-adjusted returns of both

**Table 5**

**Regional performance.** This table reports the out-of-sample performance of equal- and value-weighted long-short portfolios for emerging market subregions. All stocks are sorted into country-neutral quintile portfolios based on their predicted returns for the next month. We investigate predictions from a linear OLS model and an ensemble (*ENS*) of non-linear machine learning models (*RF*, *GBRT*, and *NN<sub>1-5</sub>*). The sorting breakpoints are based on big stocks only, which are in the top 90% of the country's aggregated market capitalization. Panel A summarizes the baseline results as presented in Table 3, and Panel B shows the result for all countries being part of emerging Americas, Panel C combines all emerging Asian countries, and Panel D reports results for emerging countries from Europe, the Middle East, and Africa. The first two rows of each panel provide the average monthly returns of the long-short quintile (Avg), corresponding *t*-statistics (*t*), the average Fama and French (2018) six-factor alphas ( $\alpha$ ), corresponding *t*-statistics ( $t_\alpha$ ), and  $R^2$ . The next two rows show spanning alphas ( $\alpha$ ), corresponding *t*-statistic ( $t_\alpha$ ), and  $R^2$  when regressing the long-short *ENS* returns on *OLS* returns and vice versa. All *t*-statistics are calculated using Newey and West (1987) adjusted standard errors with 4 lags. The sample period is from January 2002 to December 2021.

	Equal-weighted					Value-weighted				
	Avg	<i>t</i>	$\alpha$	$t_\alpha$	$R^2$	Avg	<i>t</i>	$\alpha$	$t_\alpha$	$R^2$
<b>Panel A: Emerging Markets</b>										
<i>OLS</i>	1.38	7.82	0.97	8.02	62.42	0.84	4.64	0.28	2.72	68.25
<i>ENS</i>	1.86	11.79	1.43	15.66	58.17	1.21	7.04	0.67	8.29	67.17
<i>ENS</i> ~ <i>OLS</i>			0.73	9.01	80.28			0.49	7.83	74.66
<i>OLS</i> ~ <i>ENS</i>			-0.44	-2.10	80.28			-0.22	-1.50	74.66
<b>Panel B: Americas</b>										
<i>OLS</i>	0.70	2.73	0.51	2.56	39.51	0.75	2.83	0.37	1.70	39.83
<i>ENS</i>	0.88	4.06	0.69	3.90	25.45	0.85	3.20	0.57	2.73	33.47
<i>ENS</i> ~ <i>OLS</i>			0.45	3.45	46.98			0.33	1.95	48.58
<i>OLS</i> ~ <i>ENS</i>			0.03	0.15	46.98			0.16	0.82	48.58
<b>Panel C: Asia</b>										
<i>OLS</i>	1.46	7.59	1.13	9.34	61.82	0.84	3.95	0.37	2.93	66.71
<i>ENS</i>	1.98	11.18	1.63	17.32	60.28	1.34	7.02	0.87	8.81	67.14
<i>ENS</i> ~ <i>OLS</i>			0.74	7.97	79.70			0.62	6.85	75.00
<i>OLS</i> ~ <i>ENS</i>			-0.40	-1.83	79.70			-0.33	-1.64	75.00
<b>Panel D: Europe, the Middle East and Africa</b>										
<i>OLS</i>	1.12	6.46	0.96	6.23	18.96	0.82	4.00	0.37	2.00	26.18
<i>ENS</i>	1.57	10.27	1.32	9.33	16.23	1.13	5.54	0.59	3.38	29.44
<i>ENS</i> ~ <i>OLS</i>			0.83	7.42	51.63			0.57	4.21	46.30
<i>OLS</i> ~ <i>ENS</i>			-0.11	-0.58	51.63			0.05	0.38	46.30

machine learning models. Still, it increases the value-weighted alpha of the ensemble, as presented in Panel D.

In Panel E, we utilize machine learning models, which were never trained on emerging market stock returns; instead, the models are trained on developed markets (as defined by MSCI).<sup>11</sup> Although the models were solely trained on developed markets, we surprisingly do not observe a big performance loss but actually very similar returns. Furthermore, models that allow for non-linearities and interactions (*ENS*) still significantly outperform linear models (*OLS*). This indicates that machine learning models can create significant results even if they are evaluated on data from a totally different region.

For the robustness test in Panel F, we exclude the high-turnover characteristics, namely, *Idiovol*, *LTurnover*,  $r_{2-1}$ , *SUV*, *Illiqu*, from the feature set. While the risk-adjusted returns of the *OLS* decrease by 8% and 15%, the ensemble is even more affected as the alphas are reduced by 19% and 31%. This indicates that these high-turnover features are relatively more important for more complex methods. However, even after excluding these characteristics, the long-short portfolios based on the *ENS* can span the long-short portfolios constructed based on the *OLS*, while the converse is not the case.

In Panel G, we do not train our models on a pooled sample of all countries but separately for each of the following subregions: Central and Latin America (Americas); Asia; and Europe, Middle East, and Africa (EMEA). On the one hand, this allows the models to capture potential region-specific effects. On the other hand, each model is now trained on fewer data, which might be a drawback, especially for identifying non-linearities and interactions. We document that subregional training leads to inferior return forecasts than training models on pooled data from all subregions. This finding indicates that region-specific effects play a minor role compared to more data for out-of-sample returns. Furthermore, we find that the performance decay is more pronounced for the machine learning

<sup>11</sup> For the construction of the developed market sample, we follow the same procedure as for the emerging market sample. I.e., we use country-specific constituent lists, apply static and dynamic screens as outlined in Appendix A, compute the same set of features as for emerging markets, and estimate the machine learning models in the same way as for emerging markets.

ensemble (*ENS*), i.e., indicating that more data is better for robustly identifying non-linearities and interactions.<sup>12</sup> Nevertheless, the *OLS* long-short portfolios cannot span the *ENS* long-short portfolio, but the *ENS* spans the *OLS*.

Finally, we assess if the superior performance of the machine learning return forecasts is robust across emerging market regions in Table 5. Therefore, we divide the countries of our full sample into three regions: Central and Latin America (Americas); Asia; and Europe, Middle East, and Africa (EMEA).

Overall, the results are robust for the different sub-regions Americas, Asia, and EMEA. Both *OLS* and *ENS* yield positive and significant long-short returns and alphas for both weighting schemes, but *ENS* exhibit higher returns and *t*-statistics. Furthermore, significant and positive alphas remain in the spanning regression of *ENS* on *OLS* for all sub-regions, but no positive spanning alphas remain when regressing *OLS* on *ENS*. Comparing the results across sub-regions, we find the strongest results for Asia and EMEA and a bit weaker but still highly significant results for Americas.

## 5. Understanding the sources of return predictability

The results so far provide evidence that return forecasts based on machine learning models lead to economically and statistically superior out-of-sample long-short returns compared to traditional linear models. To further understand the source of return predictability, we first investigate the performance of the two models in higher- versus lower-risk months. Second, we explore to what extent developed markets' long-short returns can explain emerging markets' long-short returns. Third, we turn to the time-series properties of the long-short machine learning portfolios over the next 36 months after portfolio formation. Fourth, we link the profitability of the machine learning models to several proxies for limits to arbitrage. Finally, we investigate the performance of an investment strategy that considers real-life investment frictions such as short-selling restrictions and transaction costs.

### 5.1. Performance in higher-risk versus lower-risk months

The profitability of return forecasts based on machine learning models may reflect risks not captured by the standard risk factors we control so far. Hence, we examine the performance of *OLS* and *ENS* forecasts during higher- versus lower-risk months. As proxies for risk, we apply whether (i) emerging markets as a whole go up or down, (ii) the rate on long-term U.S. government bonds is going up or down, (iii) the TED spread is below or above its median value, and (iv) the time-varying risk aversion index (RABex) proposed by Bekaert et al. (2022) is below or above its median value.<sup>13</sup> Splitting the sample period into up and down markets is done, for example, by Chan et al. (1998), van der Hart et al. (2005), or Asness et al. (2019). The change in the U.S. government bond rate as a proxy for risk is motivated by the substantial financial instability experienced by emerging markets during the 'taper tantrum' in 2013 when U.S. yields surprisingly surged (cf., Estrada et al., 2016). According to Frazzini and Pedersen (2014), the TED spread is a gauge of funding conditions. Lastly, Bianchi et al. (2021b) employ RABex to investigate the link between time-varying risk aversion and excess bond returns.

Table 6 summarizes the top-bottom quintile returns for *OLS* and *ENS* for the different subsamples. For both equal- and value-weighted portfolios, we observe that the performance is somewhat higher in down-market months and months with rising bond yields. However, we also document higher returns when the TED spread is low, i.e., when funding conditions are better and in months with below-median risk aversion. Nevertheless, the quintile spreads are statistically significant and positive for all subsamples, prediction models, and weighting schemes. Furthermore, the difference between the subsamples is less pronounced for *ENS* than for *OLS*, and significant and positive alphas remain in the spanning regression of *ENS* on *OLS*. At the same time, the converse is not the case. This evidence suggests that the superiority of machine learning models compared to linear models in our sample does not stem solely from higher-risk months, at least for the definitions considered here.

### 5.2. Market integration

The robustness tests in Table 4 reveal an interesting finding: models trained solely on developed markets data perform similarly to models trained on emerging markets data in predicting emerging market stock returns. This result suggests that the pricing between developed and emerging markets could be more integrated, as indicated by the results for value and momentum returns in Cakici et al. (2013) and Hanauer and Linhart (2015). If developed and emerging markets are integrated, the developed market machine learning long-short portfolio returns would be able to explain the machine learning long-short portfolio returns for emerging markets, i.e., resulting in an insignificant  $\alpha$  in the following regression for the global and regional emerging market samples:

<sup>12</sup> Table F2 in the Appendix shows that the performance decline is most pronounced for the smaller regions Americas and EMEA while smaller for Asia. Furthermore, we compare the performance of models trained on pooled data with those trained solely on local country data. We restrict this analysis to stocks from the seven countries (Chile, Indonesia, Mexico, Malaysia, Philippines, Thailand, and Turkey) that are in our sample throughout the entire sample period. The results of this analysis are presented in Table F3. Consistent with our findings for models trained on subregional data, models trained on individual country data underperform global models.

<sup>13</sup> The TED spread is defined as the difference between the LIBOR rate and the 3-month U.S. T-bill rate. The 10-year constant maturity U.S. Treasury rate (item DGS10) and the TED spread (item TEDRATE) are from the FRED database of the Federal Reserve Bank of St. Louis, and the time-varying risk aversion index RABex is from Nancy Xu's website: <https://www.nancyxu.net/risk-aversion-index>.

**Table 6**

**Higher-risk versus lower-risk periods.** This table reports the out-of-sample performance of long-short portfolios in higher- versus lower-risk months. All stocks are sorted into country-neutral quintile portfolios based on their predicted returns for the next month. The sorting breakpoints are based on big stocks only, which are in the top 90% of the country's aggregated market capitalization. Risk proxies are whether emerging markets as a whole go up or down (*Mkt*), the rate on long-term U.S. government bonds is going up or down ( $\Delta Yield$ ), whether the TED spread is below or above its median value (*TED*), and whether the time-varying risk aversion index proposed by [Bekaert et al. \(2022\)](#) (*RABex*) is below or above its median value (*RABex*). Panel A (Panel B) summarizes the results from equal-weighting (value-weighting). The first two rows of each panel provide the average monthly long-short returns and corresponding *t*-statistics. The next two rows show spanning alphas and corresponding *t*-statistic when regressing the long-short *ENS* returns on *OLS* returns and vice versa. All *t*-statistics are calculated using [Newey and West \(1987\)](#) adjusted standard errors with 4 lags. The sample period is from January 2002 to December 2021.

Model	<i>Mkt</i> <sub>up</sub>	<i>Mkt</i> <sub>down</sub>	$\Delta Yield$ <sub>up</sub>	$\Delta Yield$ <sub>down</sub>	<i>TED</i> <sub>high</sub>	<i>TED</i> <sub>low</sub>	<i>RABex</i> <sub>high</sub>	<i>RABex</i> <sub>low</sub>
<b>Panel A: Equal-Weighted</b>								
<i>OLS</i>	1.07 (3.82)	1.82 (11.81)	1.60 (7.25)	1.17 (6.20)	1.19 (3.60)	1.55 (9.93)	1.13 (3.55)	1.63 (12.48)
<i>ENS</i>	1.73 (7.01)	2.06 (14.26)	2.00 (9.52)	1.74 (9.85)	1.68 (6.07)	2.02 (12.72)	1.70 (6.15)	2.02 (14.41)
<i>ENS</i> ~ <i>OLS</i>	0.84 (7.60)	0.47 (5.43)	0.63 (6.31)	0.81 (6.60)	0.73 (7.08)	0.68 (4.95)	0.80 (8.07)	0.56 (4.25)
<i>OLS</i> ~ <i>ENS</i>	-0.64 (-2.56)	0.03 (0.23)	-0.34 (-1.46)	-0.52 (-2.13)	-0.58 (-2.38)	-0.13 (-0.81)	-0.66 (-2.93)	0.05 (0.53)
<b>Panel B: Value-Weighted</b>								
<i>OLS</i>	0.64 (2.13)	1.12 (5.51)	1.06 (4.69)	0.63 (3.31)	0.69 (2.03)	0.96 (6.49)	0.74 (2.29)	0.93 (6.30)
<i>ENS</i>	1.12 (4.12)	1.34 (5.84)	1.40 (5.53)	1.03 (6.40)	1.12 (3.55)	1.29 (7.95)	1.07 (3.65)	1.35 (8.38)
<i>ENS</i> ~ <i>OLS</i>	0.57 (6.47)	0.41 (3.88)	0.47 (4.36)	0.52 (6.37)	0.53 (6.84)	0.45 (5.03)	0.45 (4.94)	0.48 (5.27)
<i>OLS</i> ~ <i>ENS</i>	-0.37 (-2.01)	0.08 (0.54)	-0.13 (-0.71)	-0.30 (-1.78)	-0.44 (-3.35)	0.10 (0.75)	-0.27 (-1.55)	-0.00 (-0.00)

$$r_{LS,RegionEM,t,ENS} = \alpha + \beta_1 r_{LS,GlobalDev,t,ENS} + \epsilon_t \quad (10)$$

However, the results in [Table 7](#) reveal that this is not the case. All alphas remain highly statistically significant for both the equal- and value-weighted portfolios. For the equal-weighted factor, Asia has the highest alpha of 1.47% (*t*-stat 8.66), followed by EMEA with 1.24% (*t*-stat 7.39). The value-weighted factor construction yields the highest alpha for Asia with 0.98% (*t*-stat 4.92), followed by

**Table 7**

**Market integration.** This table reports summary statistics for regressions of emerging market regions' long-short returns on developed market's long-short returns. The long-short returns are based on ensemble (*ENS*) return forecasts of non-linear machine learning models (*RF*, *GBRT*, and *NN<sub>1-5</sub>*) and are separately estimated for emerging and developed markets. All stocks are sorted into country-neutral quintile portfolios based on their predicted returns for the next month. The sorting breakpoints are based on big stocks only, which are in the top 90% of a country's aggregated market capitalization. Panel A (Panel B) summarizes the results of equal-weighting (value-weighting) of the prediction-sorted portfolios based on the different regional subsets. Each Panel provides the average monthly return of the long-short quintile (*Avg*), the alphas ( $\alpha$ ), betas ( $\beta$ ), their corresponding *t*-statistics, and *R*<sup>2</sup> with respect to the developed market ensemble machine-learning factor. All *t*-statistics are calculated using [Newey and West \(1987\)](#) adjusted standard errors with 4 lags. The sample period is from January 2002 to December 2021.

	<i>Avg</i>	<i>t</i>	$\alpha$	<i>t</i> <sub><math>\alpha</math></sub>	$\beta$	<i>t</i> <sub><math>\beta</math></sub>	<i>R</i> <sup>2</sup>
<b>Panel A: Equal-weighted</b>							
<i>Global</i> <sub>EM</sub>	1.86	11.79	1.39	9.24	0.50	7.01	32.34
<i>AME</i> <sub>EM</sub>	0.88	4.06	0.35	1.73	0.56	5.57	19.44
<i>ASIA</i> <sub>EM</sub>	1.98	11.18	1.47	8.66	0.55	6.71	28.42
<i>EMEA</i> <sub>EM</sub>	1.57	10.27	1.24	7.39	0.36	4.64	12.78
<b>Panel B: Value-weighted</b>							
<i>Global</i> <sub>EM</sub>	1.21	7.04	0.89	5.73	0.49	4.95	28.17
<i>AME</i> <sub>EM</sub>	0.85	3.20	0.53	2.30	0.49	3.68	14.37
<i>ASIA</i> <sub>EM</sub>	1.34	7.02	0.98	4.92	0.54	3.95	22.47
<i>EMEA</i> <sub>EM</sub>	1.13	5.54	0.87	4.36	0.40	4.55	10.73

EMEA with an alpha of 0.87% (*t*-stat 4.36). Furthermore, the developed market long-short portfolio returns can only explain between 10% and 33% of the variation in emerging market long-short portfolio returns, which corresponds to correlations between 32% and 57%.<sup>14</sup>

Our interpretation of these results is that, although similar relationships between firm characteristics and future stock returns exist in both developed and emerging markets, the pricing of these characteristics is still not fully integrated. Furthermore, our results suggest that investors already applying machine learning strategies in developed markets may benefit from potential diversification benefits when applying such a strategy also in emerging markets.

### 5.3. Performance for longer holding periods

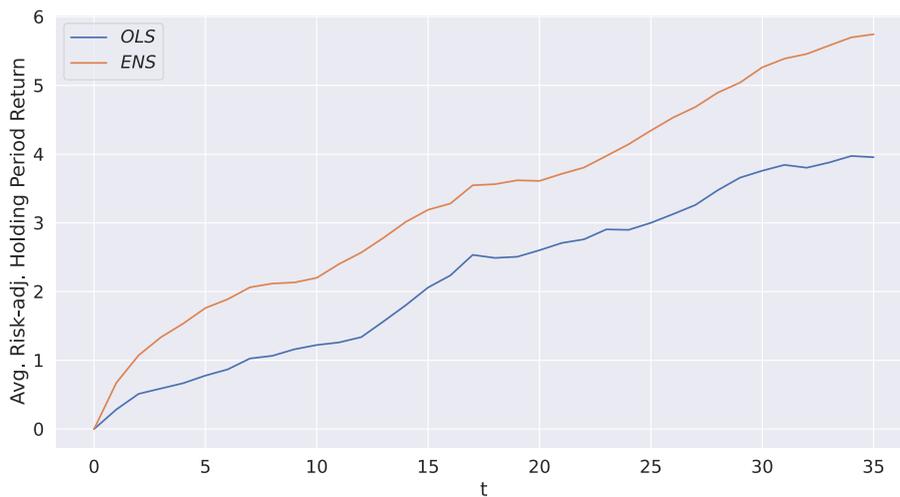
Is the profitability of the machine learning forecasts the result of temporary or permanent price changes? We analyze the long-run buy-and-hold returns following the methodology in Smajlbegovic (2019) and Hanauer et al. (2022b) to answer this question. First, we identify stocks used for constructing the long-short machine learning portfolios and calculate their value-weighted raw monthly returns for each month  $t + k$ , where  $k \in \{1, \dots, 36\}$ . Second, we run a time-series regression of the the six-factor model for each holding period month  $k$  of the machine learning long-short factor. The corresponding average six-factor alpha for month  $k$  is the intercept ( $\alpha_k$ ) of the following regression:

$$r_{t+k,ML} = \alpha_k + \sum_i^{|f|} \beta_{i,k} f_{i,t+k} + \epsilon_{t+k}, \tag{11}$$

where  $r_{t+k,ML}$  is the raw long-short return in month  $t + k$  of stocks used for construction of the long-short machine learning factor in month  $t$  and  $f_{i,t+k}$  indicates the individual factor returns of the six-factor model in month  $t + k$ :  $RMRF_{t+k}$ ,  $SMB_{t+k}$ ,  $HML_{t+k}$ ,  $RMW_{t+k}$ ,  $CMA_{t+k}$ , and  $WML_{t+k}$ . The intercept of the regression ( $\alpha_k$ ) is the alpha of the buy-and-hold strategy  $k$  months after portfolio formation, which is used to form the cumulative alpha in month  $k$  denoted as  $ACR_k$ :

$$ACR_k = \sum_{t=1}^k \alpha_t. \tag{12}$$

Fig. 6 illustrates the value-weighted cumulative six-factor alpha of both *OLS* and *ENS* over a 36-month holding period. The figure



**Fig. 6. Long-horizon performance of machine learning forecasts.** This figure shows the average cumulative risk-adjusted return of the different machine learning long-short portfolios. First, we obtain the return of the portfolio formed at the end of month  $t$  for month  $t + k$ , where  $k \in \{1, \dots, 36\}$ . Second, we run a time-series regression with the Fama and French (2018) six-factor model for the corresponding months. The regression intercept is defined as the average risk-adjusted portfolio return for the long-short portfolio in month  $t + k$ . In the final step, we compute the average holding period (cumulative) risk-adjusted return for the next  $k$  months since formation as the sum over the previous  $k$  months.

<sup>14</sup> For the construction of the developed market sample, we follow the same procedure as for the emerging market sample. I.e., we use country-specific constituent lists, apply static and dynamic screens as outlined in Appendix A, compute the same set of features as for emerging markets, and estimate the machine learning models in the same way as for emerging markets, but now on developed market data. Furthermore, we report descriptive statistics of the developed market long-short machine learning portfolio in Table F4 of the appendix. The returns for the different developed market strategies are roughly half compared to their emerging market counterparts. However, we also document that models that allow for non-linearities and interactions (*ENS*) also significantly outperform linear models (*OLS*) in developed markets.

reveals that both *OLS* and *ENS* can predict long-term returns and that their performance does not revert quickly. Together with the fact that standard risk factors cannot explain the performance of the strategies and the consistent performance over calendar time, we conclude that an underreaction explanation is more likely than an overreaction explanation. We further document that the superior performance of *ENS* compared to *OLS* is mainly driven by the first six months. Later both lines show a relatively parallel trend. This observation is unsurprising as the models are trained on one-month ahead returns and not longer periods.

#### 5.4. Limits to arbitrage

Our results thus far suggest that the high returns of the machine learning strategies in emerging markets cannot be explained by standard risk factors such as the factors of the [Fama and French \(2018\)](#) six-factor model and are consistent over time. Furthermore, the high returns do not primarily stem from higher-risk months and do not revert quickly. Therefore, a simple question arises: Why do investors not arbitrage away these abnormal returns? If limits to arbitrage hinder investors from doing that, we would expect that the predictability of the machine learning forecasts is concentrated in stocks with the highest limits to arbitrage.

To test whether the predictability of machine learning methods arises, at least in part, from such frictions, we interact the predicted returns of the machine learning models with different proxies for limits to arbitrage within a [Fama and MacBeth \(1973\)](#) regression. We additionally include both parts of the interaction term as controls as well as country dummies to account for any country effect yielding the following regression framework:

$$r_{i,t+1} - r_{f,t+1} = \alpha + \beta_1 ML_{i,t} + \beta_2 LTA_{i,t} + \beta_3 ML_{i,t} \times LTA_{i,t} + \beta_4 X_i + \epsilon_{i,t+1}, \quad (13)$$

where  $LTA_{i,t}$  denotes the cross-sectional and country-neutral standardized variable measuring the limits to arbitrage of stock  $i$  while  $ML_{i,t}$  is the predicted return based on the underlying machine learning model.

The coefficient  $\beta_3$  is most relevant for this analysis as it indicates if the predictability of the different machine learning models is increasing with higher limits to arbitrage. We include three commonly used variables that are closely related to limits to arbitrage: size as a measure of information ambiguity ([Zhang, 2006](#)), idiosyncratic volatility as a proxy for arbitrage risk ([Pontiff, 2006](#); [Stambaugh et al., 2015](#)), and [Amihud \(2002\)](#) illiquidity as a potential proxy for transaction costs. If limits to arbitrage are important for the persistence of mispricing, we expect that predictability is the strongest for smaller stocks with high idiosyncratic volatility and low liquidity. Therefore, we additionally include the average of these three variables (*COMBO*).<sup>15</sup>

The results of this analysis are reported in [Table 8](#). We first examine firm size's role in predicting future returns. Most small firms are less diversified and less fundamental information is available. In the case of fixed information acquisition costs, small firms are less attractive. The results in Columns (1) and (2) support this hypothesis. The smaller the stock, the higher the return predictability for both methods.

**Table 8**

**Limits to arbitrage.** This table reports the results of a [Fama and MacBeth \(1973\)](#) regression of next month's returns on machine learning return forecasts (ML), proxies for limits to arbitrage, and their interaction term. Each month, we conduct cross-sectional regressions of excess stock returns in month  $t + 1$  on firms' ML return forecasts, limits-to-arbitrage scores, and their interaction terms, all from the end of the previous month  $t$ . The proxies for limits to arbitrage are:  $-1 \times$  market capitalization (*SIZE*), idiosyncratic volatility (*IVOL*), Amihud illiquidity (*ILLIQ*), and a combination of the different proxies. All proxies for limits to arbitrage are ranked into the  $[-1,1]$  interval for each month and country. The  $t$ -statistics in parentheses are the corresponding [Newey and West \(1987\)](#) adjusted  $t$ -statistics with 4 lags. The sample period is from January 2002 to December 2021.

	<i>SIZE</i>		<i>IVOL</i>		<i>ILLIQ</i>		<i>COMBO</i>	
	(1)	(2)	(3)	(4)	(5)	(6)	(7)	(8)
<i>OLS</i>	0.72 (9.43)		0.72 (9.16)		0.74 (9.55)		0.74 (9.27)	
<i>ENS</i>		1.11 (13.81)		1.13 (14.31)		1.15 (14.58)		1.13 (12.88)
<i>LTA</i> $\times$ <i>ML</i>	0.27 (5.36)	0.19 (3.27)	0.48 (9.65)	0.21 (3.85)	0.01 (0.23)	-0.06 (-0.90)	0.52 (7.36)	0.27 (2.80)
<i>LTA</i>	0.15 (2.42)	0.12 (1.93)	0.06 (0.68)	0.12 (1.45)	0.21 (2.78)	0.19 (2.48)	0.27 (3.06)	0.26 (2.94)
Country dummies	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes
Adj. $R^2$ (%)	15.00	15.22	15.10	15.30	15.10	15.34	15.01	15.23
Avg. Obs	4419	4419	4419	4419	4419	4419	4419	4419

<sup>15</sup> We also report the average size, idiosyncratic volatility, illiquidity, and limit to arbitrage combination scores of the *OLS* and *ENS* portfolios in [Table F5](#) of the appendix. All proxies for limits to arbitrage are ranked into the  $[-1,1]$  interval for each month and country, where higher values indicate higher limits to arbitrage. We find relatively similar exposures for *OLS* and *ENS* portfolios. The most pronounced differences are as follows: First, the short leg of the *OLS* strategy invests on average in below-average market capitalization stocks, while the short leg of the *ENS* strategy invests in above-average market capitalization stocks. Second, the long leg of the *OLS* strategy invests on average in stocks with above-average liquidity, while the long leg of the *ENS* strategy does this to a lower extent. In combination, this leads to a below-average exposure to limits of arbitrage for the long leg of the *OLS* strategy, while the long leg of the *ENS* strategy has only a slightly below-average exposure.

In the second specification, we study how arbitrage risk affects the link between machine learning-based prediction and future stock returns. According to Pontiff (2006), arbitrageurs prefer to hold fewer stocks with higher idiosyncratic stock return volatility. Columns (3) and (4) provide empirical evidence that stocks with higher *IVOL* exhibit larger predictable returns than less volatile stocks.

Next, we test how stock illiquidity relates to our previous findings. The intuition behind this proxy is based on the tradeability of the stock. The more illiquid the stock, the slower and more costly it should be to trade on the market. However, we are not able to provide empirical evidence that the return predictability of the machine learning models is driven by transaction costs.

Finally, we combine all three limits-to-arbitrage proxies to measure their mutual influence on the effect of future return predictability. Columns (7) and (8) provide evidence that stocks associated with more substantial limits-to-arbitrage characteristics exhibit stronger predictability independent of the underlying machine learning model.

However, we also find that the higher predictability for stocks with higher limits of arbitrage is less pronounced for the machine learning ensemble *ENS* than for the linear *OLS* regression, indicating the superiority of machine learning models in emerging markets does not stem from limits to arbitrage.

### 5.5. Further investment frictions

A common feature of the results presented above is that they are based on theoretical “zero-investment” long-short portfolio returns. However, it is questionable whether these returns can be realized in practice, as short-selling constraints may prevent the implementation of long-short strategies, and transaction costs may erode the strategy’s profits. These constraints are particularly relevant for emerging markets (see, e.g., Roon et al., 2001). Therefore, in this subsection, we limit ourselves to long-only portfolios of big stocks (i.e., also remove small stocks) and consider reasonable transaction costs. To estimate transaction costs, we compute the efficient discrete generalized estimator (EDGE) of the bid-ask spread for each stock and month, recently proposed by Ardia et al. (2022). These bid-ask spread estimates vary considerably across time and stocks (cf., Fig. E3 in the appendix) and therefore provide a more sophisticated estimate than the flat 100 basis points per single-trip used in van der Hart et al. (2003) and Hanauer and Lauterbach (2019).<sup>16</sup> The transaction cost per single-trip is one-half of the estimated bid-ask spread, and we define the transaction cost of portfolio *L* as:

$$T - Cost_{L,t} = \left( \sum_{i=1}^{N_{L,t}-1} w_{i,t} - \frac{w_{i,t-1}(1+r_{i,t})}{1 + \sum_{j=1}^{N_{L,t}} w_{j,t-1}r_{j,t}} \times \frac{S_{i,t}}{2} \right), \tag{14}$$

where  $w_{i,t}$  is the weight of stock *i* at the end of month *t*,  $r_{i,t}$  is the total return of stock *i* in month *t*, and  $S_{i,t}$  is the estimated bid-ask spread. Furthermore, the net portfolio returns are defined as:

$$r_{L,t,net,ML} = r_{L,t,gross,ML} - T - Cost_{L,t}. \tag{15}$$

In the final step, we calculate the Fama and French (2018) six-factor model alpha return as:

$$r_{L,t,net,ML} - r_{f,t} = \alpha_{net} + \sum_i^{|f|} \beta_i f_{i,t,net} + \epsilon_t, \tag{16}$$

where  $f_{i,t,net}$  is the risk factor return after transaction cost.

Furthermore, we also consider trading cost mitigation rules following Novy-Marx and Velikov (2016) and Blitz et al. (2022), which are common among practitioners. Such buy/hold strategies consist of the stocks that currently belong to the top *X*% plus the stocks selected in previous months that are still among the top *Y*% of stocks. In Table 9, we compare the quintile long-only strategy (20%/20%) with the transaction-cost-mitigation strategy buying the top 10% and holding them in our portfolio as long as they belong to the top 30% (10%/30%).

Table 9 reports the strategies’ average gross excess over the market, their turnover and transaction costs, as well as the resulting net outperformance.<sup>17</sup> Limiting the investment universe to long-only portfolios of big stocks, we still see positive and significant gross outperformance for the top quintile portfolio (20%/20%) for both *OLS* and *ENS* and both weighting schemes. We observe similar gross outperformance when switching to the transaction cost mitigation strategies (10%/30%). However, the turnover and transactions are reduced by roughly 40%. This reduction in transaction costs substantially positively affects the net performance. For the equal-weighted strategies in Panel A, the net outperformance for *OLS* increases from 0.19% (*t*-stat 2.11) to 0.28% (*t*-stat 3.30). The net outperformance for *ENS* of 0.46% (*t*-stat 4.88) is also significant for the standard top quintile approach but also increases to 0.59% (*t*-stat 5.63) when applying a more efficient portfolio construction. Value-weighting the returns in Panel B leads to more challenging

<sup>16</sup> Table F6 in the appendix also provides the results for transaction cost estimates of 100 basis points per single-trip.

<sup>17</sup> Our market portfolio is a reasonable proxy for the MSCI Emerging Market Index, a popular reference index for passive investment strategies. The mean return and standard deviation for the two time-series are 1.04% and 5.88%, and 1.00% and 5.99%, respectively. Furthermore, the correlation of the two time-series is 0.97. The small differences in the returns of our market portfolio and the MSCI Emerging Market Index may be due to MSCI’s use of free float-adjusted market capitalization weighting and the inclusion of Chinese A-shares since 2018.

**Table 9**

**Further investment frictions.** This table reports the performance of different buy/hold long-only strategies before and after transaction-cost. The investment universe is limited to big stocks. We investigate predictions from a linear OLS model and an ensemble (ENS) of non-linear machine learning models (RF, GBRT, and NN<sub>1-5</sub>). Every month the portfolio consists of the stocks that currently exhibit the highest X% forecasted returns per country plus those selected in previous months whose forecasted returns have not deteriorated beyond the top Y%. The first number in the column row names represents X, while the second represents Y. We report the strategies' gross returns in excess of the market, average turnover, transaction costs, net returns in excess of the market, and net Fama and French (2018) six-factor models alphas. We estimate one-way transaction costs as one-half of a stock's bid-ask spread, estimated as in Ardia et al. (2022). All *t*-statistics are Newey and West (1987) adjusted with 4 lags. Panel A summarizes results from equal-weighting, while Panel B shows results from value-weighting. The sample period is from January 2002 to December 2021.

	OLS		ENS	
	20%/20%	10%/30%	20%/20%	10%/30%
<b>Panel A: Equal-weighted</b>				
$r_{gross}^e - Mkt$	0.49 (5.46)	0.46 (5.33)	0.78 (8.07)	0.79 (7.42)
TO (in %)	44.29	24.86	45.20	27.53
T-cost (in %)	0.31	0.18	0.32	0.20
$r_{net}^e - Mkt$	0.19 (2.11)	0.28 (3.30)	0.46 (4.88)	0.59 (5.63)
$\alpha_{net}^{FF6}$	0.29 (4.91)	0.39 (6.19)	0.55 (7.66)	0.67 (8.28)
<b>Panel B: Value-weighted</b>				
$r_{gross}^e - Mkt$	0.32 (3.39)	0.32 (3.47)	0.47 (4.88)	0.48 (4.66)
TO (in %)	44.48	22.16	45.51	23.40
T-cost (in %)	0.25	0.13	0.26	0.14
$r_{net}^e - Mkt$	0.07 (0.75)	0.19 (2.08)	0.21 (2.20)	0.34 (3.31)
$\alpha_{net}^{FF6}$	0.06 (1.26)	0.17 (3.15)	0.22 (3.91)	0.34 (5.18)

results. In this setup, the top OLS quintile yields only an insignificant net return of 0.07% (*t*-stat 0.75). Applying the trading-cost mitigation strategy increases the net returns to 0.19% (*t*-stat 2.08) for OLS and even to 0.34% (*t*-stat 3.31) for ENS. Similar results can be derived by comparing the Fama and French (2018) net alphas for which the turnover-reducing strategy for ENS exhibits again the highest net alpha of 0.34 (*t*-stat 5.18).<sup>18</sup> Therefore, we conclude that machine learning-based return forecasts can lead to significant net outperformance and net alphas, at least when efficient trading rules are applied.

## 6. Conclusion

This paper compares the out-of-sample predictive power of various machine learning models for a broad sample of 32 emerging market countries and a 20-year out-of-sample period. More specifically, we use both linear and more complex algorithms that allow for non-linearities and interactions.

We document that the different prediction algorithms identify similar characteristics. However, we also observe that tree-based methods and neural networks identify non-linearities and interactions of characteristics. Furthermore, return forecasts based on machine learning models lead to economically and statistically superior out-of-sample long-short returns compared to traditional linear models. This finding is robust to several methodological choices and for emerging market subregions.

We also find that developed market long-short returns based on machine learning forecasts derived in the same way as their emerging market counterparts cannot explain emerging market out-of-sample returns. However, models estimated solely on developed markets data can also predict emerging market stock returns. This finding indicates that similar relationships between firm characteristics and future stock returns exist for developed and emerging markets but that the pricing of these characteristics is not fully integrated between developed and emerging markets.

Furthermore, we also document that the high returns of the machine learning strategies in emerging do not primarily stem from higher-risk months and do not revert quickly, suggesting that an underreaction explanation is more likely than a risk-based explanation. Although both linear and machine learning models show higher predictability for stocks associated with higher limits to arbitrage, we also document that this effect is less pronounced for machine learning forecasts than for linear regression forecasts. This finding indicates that the superiority of machine learning models in emerging markets does not stem from limits to arbitrage. Finally, accounting for transaction costs, short-selling constraints, and limiting our investment universe to big stocks only, we document that machine learning-based return forecasts can lead to significant net outperformance over the market and net alphas, at least when

<sup>18</sup> When applying the more conservative transaction cost estimates of 100 basis points per single-trip, only the machine learning ensemble in combination with transaction cost mitigation exhibits significant net returns and alphas of 0.23% and 0.25%, respectively.

efficient trading rules are applied.

### CRedit authorship contribution statement

**Matthias X. Hanauer:** Conceptualization, Data-curation, Formal-analysis, Investigation, Methodology, Software, Validation, Visualization, Writing-review-editing. **Tobias Kalsbach:** Conceptualization, Data-curation, Formal-analysis, Investigation, Methodology, Software, Validation, Visualization, Writing-original-draft, Writing-review-editing.

### Data availability

Data will be made available on request.

## Appendix A. Filter Datastream

### A.1. Constituent lists

Datastream comprises three types of constituent lists: (1) research lists, (2) Worldscope lists, and (3) dead lists. By using dead lists, we ensure that any survivorship bias is obviated. For each country, we use the union of all available lists and eliminate any duplicates. As a result, one list remains for each country to be used in the subsequent static filter process. [Table A1](#) provides an overview of the constituent lists for emerging markets that are used in our study.

**Table A1**

**Constituent lists.** The table contains the research lists, Worldscope lists and dead lists of emerging markets countries in our sample.

Country	List	Country	List	Country	List
Argentina	DEADAR	Israel	DEADIL	Portugal	WSCOPEPT
	FARALL		WSCOPEIS		FPTALL
	WSCOPEAR		FILALL		DEADPT
Brazil	DEADBR	Jordan	DEADJO	Qatar	DEADQA
	FBRALL		FJOALL		FQAALL
	WSCOPEBR		WSCOPEJO		WSCOPEQA
Chile	DEADCL	Korea	DEADKR	Russia	DEADRU
	FCLALL		FKRALL		FRUSXALL
	WSCOPECL		WSCOPEKO		WSCPERS
China	DEADCN	Kuwait	DEADKW	Saudi Arabia	DEADSA
	FCNALL		FKWALL		FSAALL
	WSCOPECH		WSCOPEKW		WSCOPESI
Colombia	DEADCO	Malaysia	DEADMY	South Africa	DEADZA
	FCOALL		FACE		FZAALL
	WSCOPECB		FMYALL		WCOPESA
Czech Rep.	DEADCZ	Mexico	WSCOPEMY	Sri Lanka	DEADLK
	FCZALL		DEADMX		FLKALL
	WSCOPECZ		FMXALL		WSCOPECY
Egypt	DEADEC	Morocco	WSCOPEMX	Taiwan	DEADTW
	FEGALL		DEADMA		FROCOALL
	WSCOPEEY		FMAALL		FTWALL
Greece	DEADGR	Pakistan	WSCOPEMC	Thailand	WSCPETA
	FGRALL		DEADPK		DEADTH
	WSCOPEGR		FPKALL		FTHALL
Hungary	DEADHU	Peru	WSCOPEPK	Turkey	WSCPETH
	FHUALL		DEADPE		DEADTR
	WSCOPEHN		FPEALL		FTRALL
India	DEADIN	Philippines	WSCOPEPE	UAE	WSCPETK
	FINALL		DEADPH		DEADAE
	FINCONS		FPHALL		FAEALL
	FXBOMALL		WSCOPEPH		FXADSALL
	FXNSEALL		DEADPL		FXDFMALL
Indonesia	WSCOPEIN	Poland	FPLALL	WSCOPEAE	
	DEADID		FPOLCM		
	FIDALL		WSCOPEPO		
	WSCOPEID				

## A.2. Static screens

We restrict our sample to common equity stocks by applying several static screens, as shown in Table A2. Screens (1) to (7) are straightforward to apply and common in the literature.

**Table A2**

**Static screens.** The table displays the static screens applied in our study, mainly following Ince and Porter (2006), Schmidt et al. (2017) and Griffin et al. (2010). Column 3 lists the Datastream items involved (on the left of the equals sign) and the values which we set them to in the filter process (to the right of the equals sign). Column 4 indicates the source of the screens.

Nr.	Description	Datastream item(s) involved	Source
(1)	For firms with more than one security, only the one with the biggest market capitalization and liquidity is used.	MAJOR = Y	Schmidt et al. (2017)
(2)	The type of security must be equity.	TYPE = EQ	Ince and Porter (2006)
(3)	Only the primary quotations of a security are analyzed.	ISINID = P	Fong et al. (2017)
(4)	Firms are located in the respective domestic country.	GEOGN = country shortcut	Ince and Porter (2006)
(5)	Securities are listed in the respective domestic country.	GEOLN = country shortcut	Griffin et al. (2010)
(6)	Securities whose quoted currency is different to the one of the associated country are disregarded. <sup>a</sup>	PCUR = currency shortcut of the country	Griffin et al. (2010)
(7)	Securities whose ISIN country code is different to the one of the associated country are disregarded. <sup>b</sup>	GGISN = country shortcut	Annaert et al. (2013)
(8)	Securities whose name fields indicate non-common stock affiliation are disregarded.	NAME, ENAME, ECNAME	Ince and Porter (2006), Campbell et al. (2010), Griffin et al. (2010) and Karolyi et al. (2012)

<sup>a</sup> In this filter rule, the respective pre-euro currencies are also accepted for countries within the euro-zone. Moreover, in Russia 'USD' is accepted as currency, in addition to 'RUB'.

<sup>b</sup> In Hong Kong, ISIN country codes equal to 'BM' or 'KY' and in the Czech Republic ISIN country codes equal to 'CS' are also accepted.

Screen (8) relates to, among others, to work by the following: Ince and Porter (2006), Campbell et al. (2010), Griffin et al. (2010), Karolyi et al. (2012). The authors provide generic filter rules to exclude non-common equity securities from Refinitiv Datastream. We apply the identified keywords and match them with the security names provided by Datastream. A security is excluded from the sample in the event that a keyword coincides with part of the security name. The following three Datastream items store security names and are applied to the keyword filters: 'NAME', 'ENAME', and 'ECNAME'. Table A3 gives an overview of the keywords used.

**Table A3**

**Generic keyword deletions.** The table reports generic keywords searched for in the names of all stocks of all countries. If a harmful keyword is detected as part of the name of a stock, the respective stock is removed from the sample.

Non-common equity	Keywords
Duplicates	1000DUPL, DULP, DUP, DUPE, DUPL, DUPLI, DUPLICATE, XSQ, XETA
Depository receipts	ADR, GDR
Preferred stock	PF, 'PF', PFD, PREF, PREFERRED, PRF
Warrants	WARR, WARRANT, WARRANTS, WARRANT, WTS, WTS2
Debt	%, DB, DCB, DEB, DEBENTURE, DEBENTURES, DEBT
Unit trusts	.IT, .ITb, TST, INVESTMENT TRUST, RLST IT, TRUST, TRUST UNIT, TRUST UNITS, TST, TST UNIT, TST UNITS, UNIT, UNIT TRUST, UNITS, UNT, UNT TST, UT
ETFs	AMUNDI, ETF, INAV, ISHARES, JUNGE, LYXOR, X-TR
Expired securities	EXPD, EXPIRED, EXPIRY, EXPY
Miscellaneous (mainly taken from Ince and Porter (2006))	ADS, BOND, CAP.SHS, CONV, DEFER, DEP, DEPY, ELKS, FD, FUND, GW.FD, HI.YIELD, HIGH INCOME, IDX, INC.&GROWTH, INC.&GW, INDEX, LP, MIPS, MITS, MITT, MPS, NIKKEL, NOTE, OPCVM, ORTF, PARTNER, PERQS, PFC, PFCL, PINES, PRTF, PTNS, PTSH, QUIBS, QUIDS, RATE, RCPTS, REAL EST, RECEIPTS, REIT, RESPT, RETUR, RIGHTS, RST, RTN.INC, RTS, SBVTG, SCORE, SPDR, STRYPES, TOPRS, UTS, VCT, VTG.SAS, XXXXX, YIELD, YLD

In addition, Griffin et al. (2010) introduce specific keywords for individual countries. The keywords are thus applied to the security names of single countries only. For example, German security names are parsed to contain the word 'GENUSSSCHEINE', which declares the security to be a non-common equity. In Table A4, we give an overview of country-specific keyword deletions conducted in our study.

**Table A4**

**Country-specific keyword deletions.** The table reports country-specific keywords searched for in the names of all stocks of the respective countries. If a harmful keyword is detected as part of the name of a stock, the respective stock is removed from the sample.

Country	Keywords
Brazil	PN, PNA, PNB, PNC, PND, PNE, PNF, PNG, RCSA, RCTB
Greece	PR
Indonesia	FB DEAD, FOREIGN BOARD
Israel	P1, 1, 5
Korea	1P
Malaysia	'A'
Mexico	'L', 'C'
Peru	INVERSION, INVN, INV
Philippines	PDR
South Africa	N', OPTS\\, CPF\\, CUMULATIVE PREFERENCE

### A.3. Dynamic screens

For the securities remaining from the static screens above, we obtained return and market capitalization data from Datastream and accounting data from Worldscope. Several dynamic screens that are common in the literature were installed in order to account for data errors, mainly within return characteristics. The dynamic screens are shown in [Table A5](#).

**Table A5**

**Dynamic screens.** The table displays the dynamic screens applied to the data in our study, following [Ince and Porter \(2006\)](#), [Griffin et al. \(2010\)](#), [Jacobs \(2016\)](#) and [Schmidt et al. \(2017\)](#). Column 3 lists the respective Datastream items. Column 4 refers to the source of the screens.

Nr.	Description	Datastream item (s) involved	Source
(1)	We delete the zero returns at the end of the return time-series that exist because in the case of a delisting, Datastream displays stale prices from the date of delisting until the end of the respective time-series. We also delete the associated market capitalizations.	RI, MV	<a href="#">Ince and Porter (2006)</a>
(2)	We delete the associated returns and market capitalizations in case of abnormal prices (unadjusted prices > 1000000).	RI, MV, UP	The screen originally stems from <a href="#">Schmidt et al. (2017)</a> , however we employ it on unadj. price.
(3)	We delete monthly (daily) returns and the associated market capitalizations if returns exceed 990% (200%).	RI, MV	<a href="#">Griffin et al. (2010)</a> ; <a href="#">Schmidt et al., 2017</a>
(4)	We delete monthly returns and the associated market capitalizations in the case of strong return reversals, defined as $(1 + r_{t-1})(1 + r_t) - 1 < 0.5$ given that either $r_{t-1}$ or $r_t \geq 3.0$ .	RI, MV	<a href="#">Ince and Porter (2006)</a>
(5)	We delete daily returns and the associated market capitalizations in the case of strong return reversals, defined as $(1 + r_{t-1})(1 + r_t) - 1 < 0.2$ with $r_{t-1}$ or $r_t \geq 1.0$ .	RI, MV	<a href="#">Griffin et al. (2010)</a> ; <a href="#">Jacobs, 2016</a>
(6)	We delete observations of stocks that show non-zero price changes in less than 50% of the traded months in the previous 12 months.	RI, MV	<a href="#">Griffin et al. (2011)</a>
(7)	We delete observations of stocks in the lowest 3% of a country's aggregated market capitalization.	MV	<a href="#">Hanauer and Lauterbach (2019)</a>

## Appendix B. Characteristics definition

This section outlines the construction of characteristic variables that we use in the paper. For each characteristic, we give the respective Datastream and Worldscope items in parentheses, the category (past returns, investment, profitability, intangibles, value, or trading frictions) and frequency (monthly vs. yearly), plus the relevant reference. As described in Section 2, we use balance-sheet data from December in year t-1 for the stock returns from July of year t to June of year t + 1 as in [Fama and French \(1993\)](#).

**A2ME (assets-to-market), Value, Yearly.** Assets-to-market cap is the ratio of total assets (WC02999) to market capitalization as of December t-1, as in [Bhandari \(1988\)](#).

**AT (total assets), Trading Frictions, Yearly.** Total assets measured in USD (WC02999) as in [Gandhi and Lustig \(2015\)](#).

**ATO (sales-to-assets), Profitability, Yearly.** As in [Soliman \(2008\)](#), we calculate net sales (WC01001) over lagged net operating assets. Net operating assets are defined following [Hirshleifer et al. \(2004\)](#) and are explained in the construction of NOA.

**BEME (book-to-market), Value, Yearly.** Book-to-market is the ratio of book value of equity to market value of equity. We define the book value of equity as common equity (WC03501) plus deferred taxes (WC03263). If no deferred taxes are given, the book value of equity equals common equity (WC03501). The market value of equity is as of December  $t-1$ . See [Rosenberg et al. \(1985\)](#) and [Davis et al. \(2000\)](#).

**BEME<sub>m</sub>(monthly updated book-to-market), Value, Monthly.** Monthly updated book-to-market is the ratio of book value of equity to the most recent market value of equity. Book value of equity is defined as for *BEME*. The most recent market value of equity is of the end of month  $t$  to predict returns of month  $t + 1$  as in [Asness \(2011\)](#).

**Beta (market beta), Trading Frictions, Monthly.** Following [Lewellen and Nagel \(2006\)](#), we calculate beta daily as the sum of the regression coefficients of daily excess returns on the local market excess return and one lag of the local market excess return for the past 12 months. We require at least 126 observations for valid beta estimates, as in [Welch \(2020\)](#).

**C (cash-and-short-term-investment-to-assets), Value, Yearly.** The ratio of cash and short-term investments (WC02001) to total assets (WC02999), as in [Palazzo \(2012\)](#).

**CbOPtA (cash-based operating profits-to-asset), Profitability, Yearly.** As in [Ball et al. \(2016\)](#), cash-based operating profits-to-asset is operating profits converted to a cash basis divided by total assets (WC02999). Following [Ball et al. \(2015\)](#), operating profits is net sales or revenues (WC01001) minus cost of goods sold (WC01501) minus selling, general, and administrative expenses (WC01101), excluding research and development expense (WC01201). The cash-based adjustment is the year-on-year change in deferred income (WC03262), plus change in accounts payable (WC03040), plus change in accrued expenses (WC03054 + WC03069), minus change in accounts receivable (WC02051), minus change in inventory (WC02101), minus prepaid expenses (WC02140), all divided by total assets. All changes are set to zero if missing.

**CEI (composite equity issuance), Intangibles, Monthly.** Similar to [Daniel and Titman \(2006\)](#), we define composite equity issuance as the growth rate in the market capitalization not attributable to the total stock return  $R$ :  $\log(MC_{t-1}/MC_{t-13}) - R_{(t-13,t-1)}$ . To predict the returns of month  $t$ ,  $R_{(t-13,t-1)}$  is the cumulative log return (calculated via the total return index, Datastream item RI) from month  $t-13$  to month  $t-1$  and  $MC_{t-1}$  is the market capitalization (Datastream item MV) from the end of month  $t-1$ .

**CF2P (cash flow-to-price), Value, Yearly.** Cash flow to price is the ratio of net cash flow from operating activities (WC04860) to the market capitalization as of December  $t-1$ , as in [Lakonishok et al. \(1994\)](#).

**CTO (capital turnover), Profitability, Yearly.** We define capital turnover as the ratio of net sales (WC01001) to lagged total assets (WC02999), as in [Haugen and Baker \(1996\)](#).

**D2A (capital intensity), Intangibles, Yearly.** Capital intensity is the ratio of depreciation and amortization (WC01151) over total assets (WC02999), as in [Gorodnichenko and Weber \(2016\)](#).

**Debt2P (leverage), Value, Yearly.** Following [Litzenberger and Ramaswamy \(1979\)](#), debt to price is the ratio of total assets (WC02999) minus common equity (WC03501) to the market capitalization as of December  $t-1$ .

**DPI2A (ratio of change in property, plants & equipment to total assets), Investment, Yearly.** Following [Lyandres et al. \(2007\)](#), we define the changes in PP&E and inventory as the annual change in gross property, plant, and equipment (WC02301) plus the annual change in inventory (WC02101) over lagged total assets (WC02999).

**E2P (earnings-to-price), Value, Yearly.** Earnings to price is the ratio of income before extraordinary items (WC01551) to the market capitalization as of December  $t-1$ , as in [Basu \(1983\)](#).

**FC2Y (fixed costs-to-sales), Profitability, Yearly.** As in [Gorodnichenko and Weber \(2016\)](#), fixed costs to sales is the sum of selling, general and administrative expenditures (WC01101) and research and development expenses (WC01201) over net sales (WC01001).

**FreeCF (cash flow-to-book), Value, Yearly.** Following [Hou et al. \(2011\)](#), we define cash flow to book as free cash flow to book value of equity. Free cash flow is calculated as net income (WC01551) plus depreciation and amortization (WC01151) minus changes in working capital minus capital expenditure (WC04601). The book value of equity is defined in the construction of *BEME*.

**GP2A (gross profits-to-assets), Profitability, Yearly.** Gross profits-to-assets is net sales (WC01001) minus costs of goods sold (WC01051) divided by total assets (WC02999), as in [Novy-Marx \(2013\)](#).

**Idiovol (idiosyncratic volatility with respect to the Fama and French (1993) three-factor model), Trading Frictions, Monthly.** As in [Ang et al. \(2006\)](#), we define idiosyncratic volatility as the standard deviation of the residuals from a regression of excess returns on a local [Fama and French \(1993\)](#) three-factor model. We use one month of daily data and require at least fifteen non-missing observations.

**INV (investment), Investment, Yearly.** Investment is the percentage year-to-year growth rate of total assets (WC02999) following [Cooper et al. \(2008\)](#).

**LME (market capitalization), Trading Frictions, Monthly.** Size is a stock's market capitalization at the end of the previous month and measured in USD, as in [Fama and French \(1992\)](#).

**LTurnover (turnover), Trading Frictions, Monthly.** Turnover is a stock's trading volume (VO) divided by its number of outstanding shares (NOSH) during the last month, as in [Datar et al. \(1998\)](#).

**NOA (net operating assets), Investment, Yearly.** Following [Hirshleifer et al. \(2004\)](#), net operating assets are defined as the difference between operating assets and operating liabilities, scaled by lagged total assets. Operating assets are total assets (WC02999)

minus cash and short-term investments (WC02001). Operating liabilities are total assets (WC02999), minus total debt (WC03255), minus minority interest (WC03426), minus preferred stock and common equity (WC03995).

**OA (operating accruals), Intangibles, Yearly.** Following Sloan (1996), operating accruals are calculated as changes in working capital minus depreciation (WC01151) scaled by lagged total assets (WC02999). Changes in operating working capital are changes in current assets (WC02201) minus changes in cash and short-term investments (WC02001) minus changes in current liabilities (WC03101), plus changes in debt in current liabilities (WC03051) plus changes in income taxes payable (WC03063).

**OL (operating leverage), Intangibles, Yearly.** We define operating leverage as the sum of costs of goods sold (WC01051) and selling, general, and administrative expenses (WC01101) over total assets (WC02999), as in Novy-Marx (2010).

**P2P52WH (price relative to its 52-week high), Trading Frictions, Monthly.** Rel to high price is the ratio of the unadjusted stock price (UP) at the end of the previous calendar month to the past 52-weeks high, as in George and Hwang (2004).

**PCM (price-to-cost margin), Profitability, Yearly.** As in Gorodnichenko and Weber (2016) and D'Acunto et al. (2018), the price-to-cost margin is net sales (WC01001) minus costs of goods sold (WC01051), divided by net sales (WC01001).

**PM (profit margin), Profitability, Yearly.** As in Soliman (2008), we calculate the profit margin as operating income after depreciation or EBIT (WC18191) over sales (WC01001).

**Prof (gross profitability), Profitability, Yearly.** Profitability is net sales (WC01001) minus costs of goods sold (WC01051) divided by the book value of equity, following Ball et al. (2015). The book value of equity is defined in the construction of BEME.

**Q (Tobin's Q), Value, Yearly.** As in Freyberger et al. (2020), we define Tobin's Q as total assets (WC02999) plus the market capitalization as of December t-1 minus cash and short-term investments (WC02001) and minus deferred taxes (WC03263), scaled by total assets (WC02999).

**r<sub>12-2</sub>(momentum), Past Returns, Monthly.** Momentum is the cumulative return from month t-12 to t-2 as in Fama and French (1996).

**r<sub>12-7</sub>(intermediate momentum), Past Returns, Monthly.** Intermediate momentum is the cumulative return from t-12 to t-7 as in Novy-Marx (2012).

**r<sub>2-1</sub>(short-term reversal), Past Returns, Monthly.** Short-term reversal is the lagged one-month return as in Jegadeesh (1990).

**r<sub>36-13</sub>(long-term reversal), Past Returns, Monthly.** Long-term reversal is the cumulative return from t-36 to t-13 as in De Bondt and Thaler (1985).

**RNA (return on net operating assets), Profitability, Yearly.** As in Soliman (2008), we calculate the return on net operating assets as the ratio of operating income after depreciation or EBIT (WC18191) to lagged net operating assets. Net operating assets are defined following Hirshleifer et al. (2004) and explained in the construction of NOA.

**ROA (return on assets), Profitability, Yearly.** Following Balakrishnan et al. (2010), return-on-assets is the ratio of earnings before extraordinary items (WC01551) to lagged total assets (WC02999).

**ROE (return on equity), Profitability, Yearly.** Following Haugen and Baker (1996), return-on-equity are earnings before extraordinary items (WC01551) to lagged book equity. The book value of equity is defined in the construction of BEME.

**S2P (sales-to-price), Value, Yearly.** Following Lewellen (2015), sales-to-price is the ratio of net sales (WC01001) to the market capitalization as of December t-1.

**SGA2S (sales and general administrative costs to sales), Intangibles, Yearly.** As in Freyberger et al. (2020), we define SG&A as sales as the ratio of selling, general and administrative expenses (WC01101) to net sales (WC01001).

**Illiqu (Amihud (2002) illiquidity), Trading Frictions, Monthly.** We calculate illiquidity according to Amihud (2002) as the arithmetic mean of the following ratio for the past month: the daily absolute return divided by the product of the end-of-day stock price (UP) and the daily trading volume (VO).

**SUV (unexplained volume), Trading Frictions, Monthly.** Following Garfinkel (2009), standard unexplained volume is the difference between actual volume and predicted volume in the previous month. Predicted volume comes from a regression of daily volume on a constant and the absolute values of positive and negative returns. We use two months of data to estimate the model parameters (data from t-2 and t-1) and estimate the predicted volume using data from the previous month (t-1). I require at least fifteen daily observations in the previous month. Unexplained volume is standardized by the standard deviation of the residuals from the regression.

## Appendix C. Methodology

### C.1. Simple linear regression

The least complex method in our analysis and most widely used in the context of empirical asset pricing is the simple linear regression model estimated via the ordinary least squares (OLS) method. We will use it as a benchmark to compare the more complex machine learning models to it. In the case of the simple linear regression, the conditional expectations  $f^s(x)$  can be modeled using the following linear model:

$$f(x_{i,t,c}, \theta) = \theta^T x_{i,t,c}, \quad (17)$$

where  $\theta, \theta^T = (\theta_1, \theta_2, \dots, \theta_p) \in \mathbb{R}^p$ , is the column vector of coefficients that can be estimated with OLS by minimizing the loss function:

$$L_{MSE}(\theta) = \frac{1}{NT} \sum_{i=1}^N \sum_{t=1}^T (r_{i,t+1,c}^{abn} - f(x_{i,t,c}, \theta))^2, \quad (18)$$

which is also known as the Mean Squared Error (MSE). The OLS has the big advantage that it does not require any hyperparameter input from the user. Further, by minimizing the loss function  $L_{MSE}$  a unique analytical solution can be extracted, which is easy to interpret as the coefficients,  $\theta$  directly describe how a change in the stock characteristics affects the expected return. Additionally, if the number of observations in the underlying dataset is larger than the number of coefficients that need to be estimated, the OLS yields an efficient and unbiased estimator according to Wooldridge (2001). But if the number of characteristics approaches the number of observations in the dataset, the OLS has issues distinguishing between signal and noise. While the signal is the portion we can understand, model, and predict, noise consists of the unpredictable component of price movements. In the case of a small sample or a large number of characteristics, the OLS starts with over-fitting the coefficients to noise rather than extracting the signal. This is of particular importance in the field of asset pricing, which empirically relies on a low signal-to-noise ratio. This overfitting yields a higher in-sample performance but a poor out-of-sample performance. Further, multicollinearity between the different characteristics can lead to a fallacious interpretation of test statistics as well as misleading coefficients. Lastly, the OLS does not model or evaluate any non-linearities of the characteristics nor any potential interactions between them. Any non-linearity would have to be imputed by the user.

### C.2. Regularized regression

To avoid overfitting in the case of empirical asset pricing, the user could increase the training sample, reduce the number of characteristics used to predict future returns, or utilize regularized regression techniques that identify which characteristics are informative and omits those that are not. Classical regularized regression techniques include ridge regression, lasso regression, and elastic net. To limit the number of machine learning methods, we concentrate on the elastic net, which is a combination of ridge and lasso regression. While the different regularized regression models have the same linear functional form as the simple linear regression, they differ with respect to the loss function by adding a penalty term ( $\phi_{ENet}(\theta, \lambda, \alpha)$ ) to it:

$$L_{ENet}(\theta, \lambda, \alpha) = L_{MSE}(\theta) + \phi_{ENet}(\theta, \lambda, \alpha). \quad (19)$$

This penalty term reduces the model's in-sample performance and increases its out-of-sample stability by shrinking the coefficients of noisy characteristics, improving the signal-to-noise ratio. The penalty function of the elastic net is defined as:

$$\phi_{ENet}(\theta, \lambda, \alpha) = (1 - \alpha)\lambda \sum_{j=1}^P |\theta_j| + \frac{1}{2}\alpha\lambda \sum_{j=1}^P \theta_j^2, \quad (20)$$

where  $\lambda, \lambda \in \mathbb{R}^+$  defines the magnitude of shrinkage and  $\alpha, \alpha \in \{0, \dots, 1\}$  which determines the relative weight between the two penalty components of the ridge and lasso regression. In the case of  $\lambda = 0$  the regularized regression models yield a simple linear regression model. The coefficients are shrunk towards zero by setting  $\lambda > 0$ . As these two hyperparameters have to be set by the user, we utilize our validation sample to find the optimal in-sample  $\lambda$  and  $\alpha$  in the first run. We determine the optimal  $\theta$  in the second run using the full training and validation sample.

### C.3. Tree-based regression

Tree-based models represent the first non-parametric regression model as their structure is decided by the training data. For our return prediction, we will utilize two tree-based methods: the random forest, as well as the gradient-boosted regression tree. Compared to the linear methods, one advantage of these tree methods is that the user does not have to manually add any potential non-linearities or interactions to the data as the tree methods build these by construction.

Regression trees follow the idea of sequentially partitioning the underlying data into groups that behave similarly to each other based on a selected characteristic with regard to the future return. By sequentially separating the data, the tree "grows" and new "branches" are created each time the data is split into new groups. The tree can grow to a depth of  $D$  based on the user input. At each new branch, the characteristic is picked that causes the biggest separation in the data based on an optimized cut-off value.<sup>19</sup> As soon as the data cannot be split into subgroups or the depth  $D$  is reached, a "leaf" is created. In asset pricing, the tree yields a return that is clustered by the underlying characteristics.

The following equation describes a tree with a depth of  $D$  and  $K$  leaves:

<sup>19</sup> In our case, for each separation, the characteristic is selected that minimizes the MSE.

$$\begin{aligned}
 f(x_{i,t,c}, \theta, D, K) &= \sum_{k=1}^K \theta_k \mathbb{1}_{\{x_{i,t,c} \in C_k(D)\}} \\
 \theta_k &= \frac{1}{N_k} \sum_{x_{i,t,c} \in C_k(D)} r_{i,t+1,c}^{\text{abn}}
 \end{aligned}
 \tag{21}$$

where  $D$  is the depth of the tree measured as the maximum number of separations following the longest branch,  $C_k(D)$  indicates the  $k$ -th separation of the characteristics,  $\theta_k$  is average abnormal return within the partition, and  $\mathbb{1}_{\{x_{i,t,c} \in C_k(D)\}}$  indicates if  $x_{i,t,c}$  is part of  $C_k(D)$ . Following this methodology, a tree of depth  $D$  can capture up to  $D - 1$  interactions. To avoid overfitting, the tree must be regularized. We follow two different approaches in our analysis.

The first regularization approach uses bootstrap aggregation, or “bagging”, developed by Breiman (2001). In this approach, each of the  $T$  trees starts with a share of  $B$  bootstrap samples from the data and fits an individual regression tree to the bootstrapped data. Afterward, the forecasts from the individual trees are averaged. This reduces the variation in the prediction and stabilizes the prediction performance. In the case of the random forest, the trees additionally use random subsets  $R$  of characteristics to grow the branches. This reduces the impact of certain dominant return characteristics and creates de-correlated trees.

The second regularization approach is “boosting.” It starts by training a weak and shallow regression tree on the full training data. In the next step, a second regression tree with the same depth  $D$  is trained on the residuals of the first tree. The prediction of these two trees is then averaged while the contribution of the second tree is shrunk by a factor  $LR$  (learning rate),  $LR \in (0, 1)$  to avoid the model overfitting the residuals. At each new step  $b$ , till the model reaches a total of  $B$  trees, a new shallow tree is fitted to the residual, which is based on the  $b - 1$ -th model and added to it with a shrinkage weight of  $LR$ .

Both regression trees share the two main hyperparameters: the number of trees in the forest  $T, T \in \mathbb{Z}^+$  and the maximum depth  $D, D \in \mathbb{Z}^+$ . While the random forest additionally requires the share of the bootstrapped samples  $B, 0 > B \leq 1$ , the gradient-boosted regression tree requires a certain learning rate  $LR, 0 > LR \leq 1$ . These hyperparameters are optimized through the validation step. Additionally, we can provide the share  $R, 0 > R \leq 1$ , of randomly selected characteristics that are used in each tree of the random forest.

#### C.4. Neural networks

Neural networks are another highly flexible but opposed to the regression trees, a parametric model. While these models can have various forms, we focus on the standard structure of a feed-forward neural network. A feed-forward neural network consists of an “input” layer of input characteristics and the intercept, at least one “hidden” layer comprising activation functions, and an “output” layer that aggregates the outcome of the last hidden layer into a return prediction.

A feedforward neural network consists of several subsequent layers  $l, l = 0, 1, \dots, L$ , one input layer ( $l = 0$ ),  $L - 1$  hidden layers ( $l = 1, 2, \dots, L - 1$ ) and one output layer  $l = L$ . Each layer  $l$  contains  $n^l$  nodes. In the case of the input layer, the number of nodes is equal to the number of characteristics, including an intercept, while the output layer contains due to the regression setting one node. In the case of the hidden layer, we consider an architecture of up to five hidden layers while the first hidden layer contains 32 nodes and each additional hidden layer divides the number of nodes by two compared to the previous layer following the geometric pyramid rule according to Masters (1993). This results in the following number of nodes per layer:

$$\begin{aligned}
 n^0 &= p + 1, \\
 n^1 &= 32, \\
 n^l &= \frac{n^{l-1}}{2} \forall l \in \{2, \dots, L - 1\}, \\
 n^L &= 1.
 \end{aligned}
 \tag{22}$$

Each of the nodes in the hidden layer contains an activation function. In our case we follow Gu et al. (2020) and Leippold et al. (2022) and choose the rectified linear unit defined as:

$$\text{ReLU}(x) = \max(0, x),
 \tag{23}$$

As in De Nard et al. (2022), we adopt the Adam optimization algorithm (Kingma and Ba, 2014), early stopping, batch normalization (Ioffe and Szegedy, 2015), ten ensembles with individual seeds (Hansen and Salamon, 1990; Dietterich, 2000) and dropout (Srivastava et al., 2014) when training our models.

#### C.5. Hyperparameters

We will use the following hyperparameters based on the hyperparameter range in Gu et al. (2020), Tobek and Hronec (2020),

Drobetz and Otto (2021), Leippold et al. (2022):

- Elastic net
  - $\lambda$ :  $[1 \times 10^{-5}, 2 \times 10^{-5}, \dots, 1 \times 10^{-2}]$
  - $\alpha$ :  $[0, 0.01, \dots, 1]$
- Random forest
  - $R$ :  $[0.01, 0.02, \dots, 1]$
  - $B$ : 1
  - $T$ :  $[100, 102, \dots, 600]$
  - $D$ :  $[1, 2, \dots, 8]$
- Gradient-boosted regression tree
  - $LR$ :  $[0.01, 0.02, \dots, 0.1]$
  - $T$ :  $[50, 52, \dots, 500]$
  - $D$ :  $[1, 2, \dots, 8]$
- Neural networks
  - $l_1$ :  $[0.00001, \dots, 0.001]$
  - $LR$ :  $[0.001, 0.1]$
  - Batch Size: 10000
  - Epochs: 100

#### Appendix D. Factor construction

This section outlines the construction of the factors of the Fama and French (2018) six-factor model, using the same stock sample as for the machine learning portfolios described in Section 2.1.

The market factor, RMRF, is the value-weighted return of all stocks minus the risk-free rate. The remaining factors are constructed using a  $2 \times 3$  sorting approach commonly employed for international markets (Fama and French, 2012; Fama and French, 2017). The portfolio breakpoints for the value, profitability, investment, and momentum factors are the 30% and 70% percentiles of the underlying characteristic of the big-stock sample for each country. In the case of the value factor, we use the book-to-market ratio to form Growth (G), Neutral (N), and Value (V) portfolios. For profitability, we use cash-based operating profitability to sort the stocks into the extreme portfolios Weak (W) and Robust (R). We use asset growth for the investment factor, which yields Conservative (C) and Aggressive (A) portfolios. For the momentum factor, we sort stocks into the Winner (W) and Loser (L) portfolios based on a stock's momentum. Finally, we classify stocks into the two size groups big (B) and small (S), as described in Section 2.1. The final factor calculation is based on the intersection of the different portfolios, while the portfolio returns are value-weighted,

$$\begin{aligned}
 SMB &= (SV + SN + SG)/3 - (BV + BN + BG)/3, \\
 HML &= (BV + SV)/2 - (BG + SG)/2, \\
 RMW &= (BR + SR)/2 - (BW + SW)/2, \\
 CMA &= (BC + SC)/2 - (BA + SA)/2, \\
 WML &= (BW + SW)/2 - (BL + SL)/2.
 \end{aligned} \tag{24}$$

Table D1 presents summary statistics for the monthly factor returns.

**Table D1**

**Summary statistics for the factors of the Fama and French (2018) six-factor model.** This table presents the average monthly return, standard deviation, and corresponding t-statistic for the following set of factors: the market (RMRF), size (SMB, small minus big), value (HML, high minus low), profitability (RMW, robust minus weak based on cash-based operating profitability), investment (CMA, conservative minus aggressive), and momentum (WML, winners minus losers). The t-statistics are Newey-West adjusted with 4 lags. The sample period for the analysis is January 2002 to December 2021.

	RMRF	SMB	HML	CMA	RMW	WML
mean	0.94	–0.0	0.52	0.07	0.49	0.79
std. dev.	5.88	1.36	1.89	2.24	2.65	4.16
t-stat	2.05	–0.05	3.40	0.38	2.45	2.54

#### Appendix E. Figures

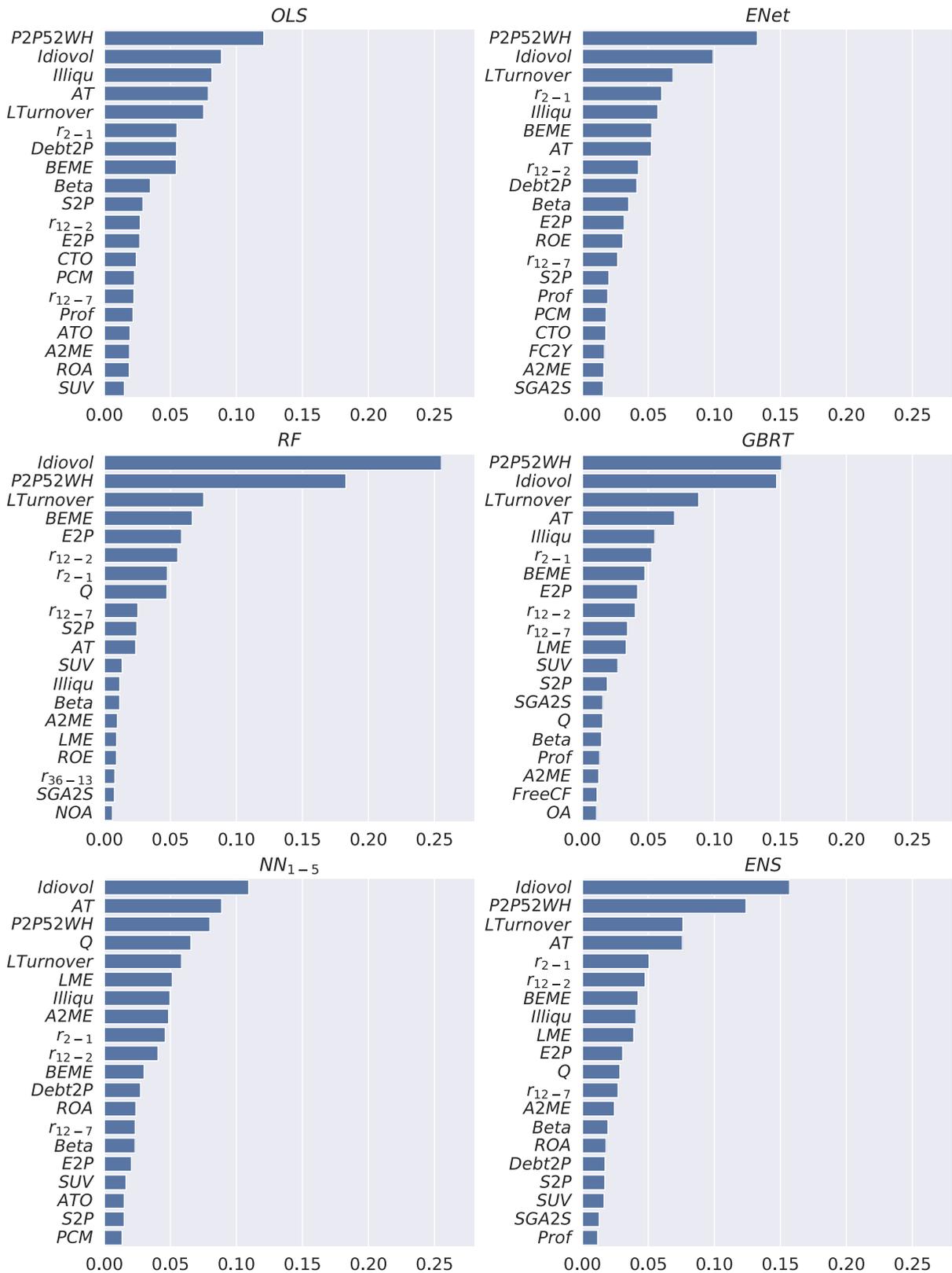
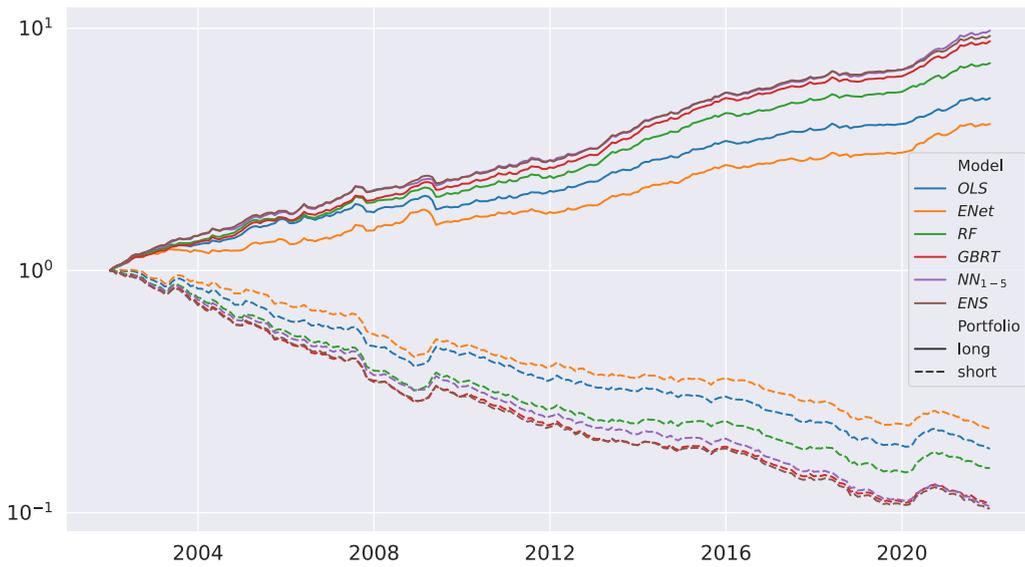
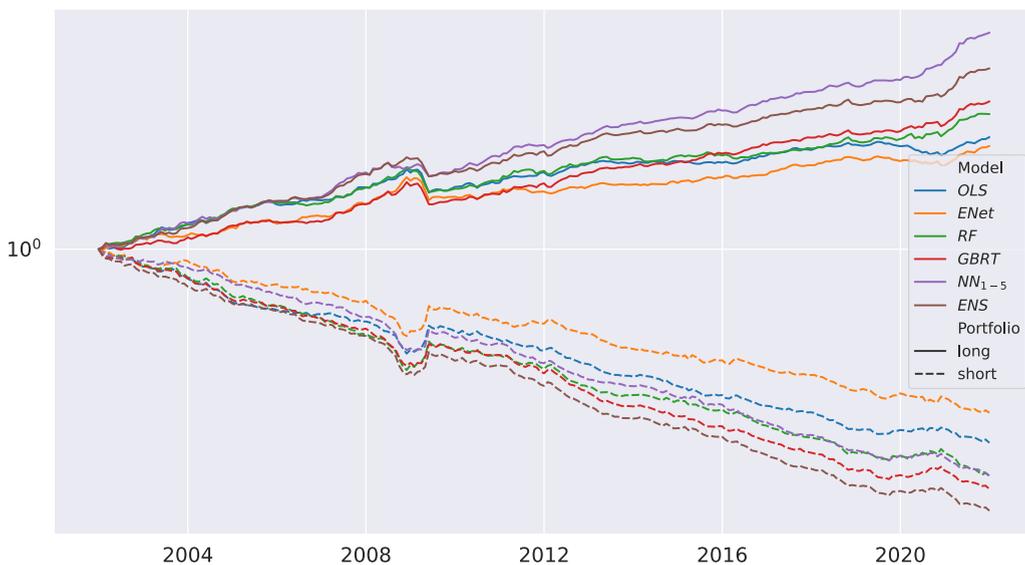


Fig. E1. Variable importance by model. This figure shows the importance for Individual characteristics in each model. Characteristics importance is an average over all training samples. Variable importance within each model is normalized to sum to one.

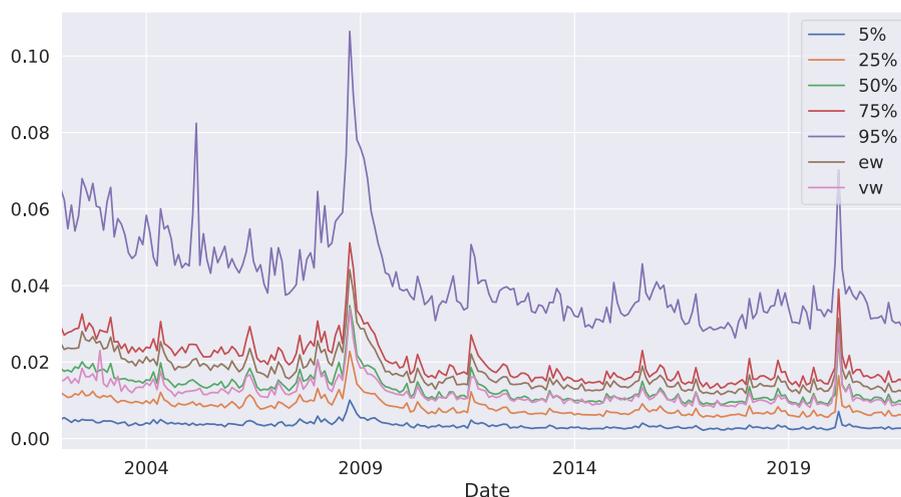
Panel A: Equal-weighted



Panel B: Value-weighted



**Fig. E2. Cumulative return of machine learning portfolios.** The figure shows the cumulative log returns in excess of the market of portfolios sorted on out-of-sample machine learning return forecasts. The solid and dashed lines represent long (top quintile) and short (bottom quintile) portfolios, respectively. In Panel A equal-weighted cumulative log returns are shown while in Panel B the long and short portfolios are value-weighted. The sample period is from January 2002 to December 2021.



**Fig. E3. Estimated bid-ask spreads based on the EDGE estimator.** This figure shows the cross-sectional distribution of estimated bid-ask spreads for big stocks in emerging markets. Thereby, big stocks are defined as the biggest stocks, which together account for 90% of a country's aggregated market capitalization. For each stock and month, we compute the efficient discrete generalized estimator (EDGE) of the bid-ask spread, proposed in [Ardia et al. \(2022\)](#). The estimators are based on daily prices using a monthly estimation window. Following [Novy-Marx and Velikov \(2016\)](#), we replace zero estimates with the non-zero estimate of the stock of the same country with the shortest Euclidean distance in size and characteristic volatility rank space. The sample period is from January 2002 to December 2021.

## Appendix F. Tables

**Table F1**

**Detail performance of the machine learning portfolios.** This table reports the out-of-sample performance of the different machine learning quintile portfolios. Stocks are sorted into country-neutral quintiles based on their predicted returns for the next month. The sorting breakpoints are based on big stocks only, which are in the top 90% of a country's aggregated market capitalization. Each Panel provides the predicted monthly returns (Pred), the average monthly excess returns (Avg), corresponding  $t$ -statistics ( $t$ ), the [Fama and French \(2018\)](#) six-factor model alpha ( $\alpha$ ), and corresponding  $t$ -statistics. All  $t$ -statistics are calculated using [Newey and West \(1987\)](#) adjusted standard errors with 4 lags. The sample period is from January 2002 to December 2021.

	Equal-weighted					Value-weighted				
	Pred	Avg	$t$	$\alpha$	$t_\alpha$	Pred	Avg	$t$	$\alpha$	$t_\alpha$
<b>Panel A: OLS</b>										
Low (L)	-1.07	0.26	0.50	-0.39	-3.43	-0.99	0.43	0.83	-0.23	-3.40
2	-0.38	0.90	1.85	0.11	1.16	-0.37	0.82	1.70	-0.00	-0.05
3	0.00	1.15	2.51	0.26	3.08	0.01	0.95	2.09	0.09	2.55
4	0.36	1.34	3.03	0.41	4.91	0.37	1.13	2.55	0.17	3.45
High (H)	0.86	1.64	3.73	0.58	6.39	0.87	1.25	2.82	0.06	1.05
H-L	1.93	1.38	7.76	0.97	8.02	1.85	0.83	4.57	0.28	2.72
<b>Panel B: ENet</b>										
Low (L)	-1.07	0.34	0.65	-0.33	-3.05	-0.99	0.51	0.98	-0.18	-2.33
2	-0.37	0.92	1.86	0.12	1.32	-0.36	0.81	1.70	0.02	0.32
3	0.02	1.16	2.48	0.27	3.19	0.03	0.96	2.06	0.09	1.95
4	0.39	1.32	2.97	0.38	4.52	0.40	1.09	2.43	0.11	2.21
High (H)	0.90	1.53	3.61	0.50	4.99	0.91	1.23	2.86	0.09	1.56

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Table F1 (continued)

	Equal-weighted					Value-weighted				
	Pred	Avg	<i>t</i>	$\alpha$	$t_\alpha$	Pred	Avg	<i>t</i>	$\alpha$	$t_\alpha$
H-L	1.97	1.20	6.79	0.83	6.94	1.90	0.72	3.95	0.27	2.28
<b>Panel C: RF</b>										
Low (L)	-0.79	0.18	0.35	-0.47	-4.05	-0.69	0.34	0.66	-0.30	-4.37
2	-0.21	0.82	1.74	0.07	0.81	-0.21	0.86	1.82	0.09	1.77
3	0.09	1.12	2.41	0.23	2.62	0.09	0.94	2.06	0.03	0.61
4	0.37	1.35	2.92	0.40	5.16	0.37	1.15	2.51	0.09	1.45
High (H)	0.71	1.78	3.96	0.72	8.34	0.70	1.32	2.99	0.17	3.38
H-L	1.50	1.60	9.29	1.19	14.10	1.39	0.99	5.24	0.47	5.24
<b>Panel D: GBRT</b>										
Low (L)	-0.87	0.04	0.08	-0.59	-5.24	-0.74	0.30	0.58	-0.38	-5.57
2	-0.18	0.87	1.82	0.11	1.21	-0.17	0.81	1.71	0.09	1.68
3	0.13	1.13	2.42	0.25	3.25	0.13	0.94	2.04	0.03	0.73
4	0.43	1.36	2.98	0.39	4.76	0.42	1.12	2.47	0.11	1.62
High (H)	0.93	1.86	4.12	0.81	8.77	0.86	1.35	3.09	0.19	4.06
H-L	1.80	1.82	11.49	1.40	15.65	1.61	1.05	6.06	0.57	6.73
<b>Panel E: NN<sub>1</sub></b>										
Low (L)	-1.27	0.03	0.06	-0.62	-5.23	-1.05	0.41	0.81	-0.26	-2.74
2	-0.28	0.80	1.71	0.04	0.54	-0.26	0.76	1.68	-0.01	-0.28
3	0.16	1.07	2.31	0.21	2.60	0.16	0.94	2.08	0.05	1.15
4	0.58	1.34	2.93	0.40	4.92	0.57	1.07	2.35	0.05	0.78
High (H)	1.34	1.91	4.09	0.86	8.93	1.25	1.45	3.09	0.30	5.86
H-L	2.61	1.88	13.92	1.47	15.67	2.30	1.04	6.94	0.57	4.83
<b>Panel F: NN<sub>2</sub></b>										
Low (L)	-1.27	-0.01	-0.03	-0.68	-6.33	-1.01	0.37	0.74	-0.35	-5.42
2	-0.23	0.82	1.74	0.06	0.70	-0.21	0.80	1.72	0.03	0.52
3	0.17	1.01	2.18	0.16	2.10	0.18	0.93	2.08	0.03	0.65
4	0.56	1.35	3.06	0.42	4.76	0.55	1.08	2.42	0.06	0.87
High (H)	1.32	1.90	3.99	0.86	8.87	1.21	1.49	3.08	0.36	6.84
H-L	2.60	1.91	15.67	1.55	19.02	2.21	1.11	9.40	0.71	9.16
<b>Panel G: NN<sub>3</sub></b>										
Low (L)	-1.20	0.02	0.05	-0.66	-5.87	-0.94	0.38	0.74	-0.35	-4.59
2	-0.18	0.82	1.75	0.06	0.74	-0.16	0.73	1.60	-0.01	-0.34
3	0.16	1.08	2.32	0.24	3.47	0.17	0.97	2.17	0.09	2.29
4	0.51	1.31	2.92	0.39	4.42	0.50	1.08	2.36	0.06	1.13
High (H)	1.22	1.86	3.99	0.83	8.34	1.11	1.49	3.17	0.32	5.34
H-L	2.41	1.84	14.72	1.49	16.93	2.04	1.12	7.85	0.66	6.55
<b>Panel H: NN<sub>4</sub></b>										
Low (L)	-1.14	0.04	0.07	-0.63	-5.57	-0.90	0.32	0.62	-0.38	-5.18
2	-0.17	0.79	1.67	0.03	0.34	-0.14	0.73	1.59	-0.02	-0.50
3	0.16	1.09	2.37	0.26	3.34	0.17	0.95	2.14	0.08	1.87
4	0.48	1.31	2.91	0.35	4.39	0.47	1.11	2.43	0.05	0.81
High (H)	1.15	1.89	4.07	0.86	8.26	1.04	1.52	3.25	0.35	6.35
H-L	2.29	1.86	13.75	1.48	16.72	1.94	1.20	8.30	0.73	7.76
<b>Panel I: NN<sub>5</sub></b>										
Low (L)	-1.12	0.04	0.08	-0.62	-5.43	-0.90	0.36	0.68	-0.37	-4.83
2	-0.17	0.82	1.74	0.04	0.46	-0.15	0.71	1.56	0.01	0.29
3	0.18	1.10	2.37	0.27	3.68	0.18	0.91	1.97	0.02	0.50
4	0.50	1.32	2.96	0.38	4.59	0.49	1.13	2.54	0.08	1.47
High (H)	1.14	1.88	4.04	0.82	8.38	1.04	1.52	3.26	0.34	7.61
H-L	2.26	1.84	13.42	1.44	15.79	1.93	1.17	8.11	0.71	8.21

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Table F1 (continued)

	Equal-weighted					Value-weighted				
	Pred	Avg	$t$	$\alpha$	$t_\alpha$	Pred	Avg	$t$	$\alpha$	$t_\alpha$
<b>Panel J: <math>NN_{1-5}</math></b>										
Low (L)	-1.13	0.03	0.06	-0.62	-5.41	-0.91	0.34	0.65	-0.36	-4.82
2	-0.19	0.77	1.65	-0.00	-0.00	-0.16	0.73	1.62	0.02	0.42
3	0.17	1.10	2.38	0.27	3.57	0.17	0.95	2.10	0.06	1.35
4	0.51	1.31	2.92	0.39	4.50	0.50	1.07	2.39	0.03	0.54
High (H)	1.17	1.90	4.05	0.84	8.70	1.07	1.54	3.22	0.36	7.02
H-L	2.30	1.87	13.42	1.46	15.81	1.97	1.21	8.48	0.72	8.26
<b>Panel K: ENS</b>										
Low (L)	-0.85	0.02	0.04	-0.61	-5.32	-0.71	0.24	0.47	-0.42	-5.89
2	-0.17	0.85	1.78	0.09	0.99	-0.16	0.80	1.71	0.09	1.85
3	0.13	1.13	2.46	0.29	3.44	0.13	0.90	1.98	0.01	0.29
4	0.41	1.38	3.02	0.40	5.17	0.41	1.15	2.59	0.12	1.95
High (H)	0.87	1.88	4.12	0.82	9.02	0.81	1.45	3.17	0.25	5.82
H-L	1.71	1.86	11.73	1.43	15.66	1.52	1.20	6.97	0.67	8.29
<b>Panel L: <math>\mu_{\text{sign}(c)}</math></b>										
Low (L)	-0.20	0.22	0.42	-0.41	-3.16	-0.19	0.48	0.93	-0.22	-3.42
2	-0.08	0.79	1.57	0.00	0.02	-0.08	0.90	1.84	0.10	2.10
3	-0.01	1.05	2.20	0.18	2.19	-0.01	0.96	2.10	0.08	1.75
4	0.06	1.25	2.77	0.32	3.72	0.06	1.08	2.45	0.05	1.06
High (H)	0.15	1.56	3.67	0.57	7.27	0.15	1.24	2.91	0.11	2.18
H-L	0.35	1.34	6.48	0.98	7.26	0.34	0.77	4.16	0.32	3.32

Table F2

**Robustness - Models trained on subregional data.** This table reports the out-of-sample performance of equal-weighted and value-weighted long-short portfolios sorted on forecasts derived from models trained on subregional data. All stocks are sorted into country-neutral portfolios based on their predicted returns for the next month. The sorting breakpoints are based on big stocks only, which are in the top 90% of the country's aggregated market capitalization. Panel A shows the results pooled emerging markets, Panel B for all countries being part of emerging Americas, Panel C combines all emerging Asian countries, and Panel D reports results for emerging countries from Europe, the Middle East, and Africa. The first two rows of each panel provide the average monthly return of the long-short quintile (Avg), corresponding  $t$ -statistics ( $t$ ), the average Fama and French (2018) six-factor alpha ( $\alpha$ ), corresponding  $t$ -statistics ( $t_\alpha$ ), and  $R^2$ . The next two rows show spanning alpha ( $\alpha$ ), corresponding  $t$ -statistic ( $t_\alpha$ ), and  $R^2$  when regressing the long-short ENS returns on OLS returns and vice versa. All  $t$ -statistics are calculated using Newey and West (1987) adjusted standard errors with 4 lags. The sample period is from January 2002 to December 2021.

	Equal-weighted					Value-weighted				
	Avg	$t$	$\alpha$	$t_\alpha$	$R^2$	Avg	$t$	$\alpha$	$t_\alpha$	$R^2$
<b>Panel A: Emerging Markets</b>										
OLS	1.09	6.29	0.77	6.23	53.92	0.78	4.42	0.29	2.57	56.92
ENS	1.35	8.88	0.97	9.75	57.22	0.97	5.95	0.44	4.59	58.10
ENS ~ OLS			0.49	7.78	77.56			0.28	4.46	75.77
OLS ~ ENS			-0.23	-1.76	77.56			-0.06	-0.51	75.77
<b>Panel B: Americas</b>										
OLS	0.76	3.23	0.36	1.82	46.02	0.60	2.67	0.04	0.22	49.64
ENS	0.70	3.37	0.41	2.52	36.93	0.64	3.12	0.20	1.14	35.55
ENS ~ OLS			0.19	1.56	56.53			0.22	1.66	49.56
OLS ~ ENS			0.17	1.02	56.53			0.14	0.83	49.56
<b>Panel C: Asia</b>										
OLS	1.43	7.73	1.12	9.62	60.08	0.81	4.03	0.41	3.22	63.14
ENS	1.95	11.16	1.61	17.97	60.73	1.27	6.62	0.83	8.24	67.10
ENS ~ OLS			0.65	8.64	83.54			0.52	5.07	71.55
OLS ~ ENS			-0.36	-2.10	83.54			-0.17	-1.11	71.55

(continued on next page)

Table F2 (continued)

	Equal-weighted					Value-weighted				
	Avg	$t$	$\alpha$	$t_\alpha$	$R^2$	Avg	$t$	$\alpha$	$t_\alpha$	$R^2$
<b>Panel D: Europe, the Middle East and Africa</b>										
<i>OLS</i>	1.06	5.97	0.91	5.67	19.35	0.92	4.43	0.47	2.39	26.25
<i>ENS</i>	1.39	8.11	1.06	6.59	20.58	0.99	4.75	0.34	1.97	37.32
<i>ENS ~ OLS</i>			0.61	5.47	56.62			0.25	2.01	59.45
<i>OLS ~ ENS</i>			-0.01	-0.06	56.62			0.19	1.45	59.45

Table F3

**Robustness - Models trained on pooled versus individual countries.** This table reports the out-of-sample performance of value-weighted long-short portfolios sorted on local and pooled model forecasts. Local model forecasts are based on machine learning models trained separately for each country on local country data. In contrast, the pooled model forecasts are from our baseline machine learning models trained on pooled emerging market data. Both local and pooled strategies include only stocks from the following seven countries that are in our sample throughout the entire sample period: Chile, Indonesia, Mexico, Malaysia, Philippines, Thailand, and Turkey. All stocks are sorted into country-neutral quintile portfolios based on predicted returns for the next month. The sorting breakpoints are based on big stocks only. The first two rows of each panel provide the average monthly excess returns (Avg), corresponding  $t$ -statistics ( $t$ ), the average Fama and French (2018) six-factor alphas ( $\alpha$ ), corresponding  $t$ -statistics ( $t_\alpha$ ), and  $R^2$ . The next two rows show spanning alphas ( $\alpha$ ), corresponding  $t$ -statistics ( $t_\alpha$ ), and  $R^2$  when regressing the long-short local returns on pooled returns and vice versa. All  $t$ -statistics are calculated using Newey and West (1987) adjusted standard errors with 4 lags. The sample period is from January 2002 to December 2021.

	<i>OLS</i>					<i>ENS</i>				
	Avg	$t$	$\alpha$	$t_\alpha$	$R^2$	Avg	$t$	$\alpha$	$t_\alpha$	$R^2$
<i>local</i>	0.61	3.68	0.41	2.83	17.13	0.92	5.93	0.63	4.70	23.16
<i>pooled</i>	0.80	4.37	0.41	2.26	34.41	1.18	6.67	0.81	5.66	27.08
<i>local ~ pooled</i>			0.17	1.42	38.67			0.25	1.83	40.26
<i>pooled ~ local</i>			0.37	2.15	38.67			0.53	3.15	40.26

Table F4

**Performance of machine learning portfolios for developed markets.** This table reports the out-of-sample performance of long-short portfolios in developed markets. Stocks are sorted into country-neutral quintile portfolios based on the predicted returns from machine learning models trained with developed market data. The sorting breakpoints are based on big stocks only. The first two rows of each panel provide the average monthly returns of the long-short quintiles (Avg), corresponding  $t$ -statistics ( $t$ ), the average Fama and French (2018) six-factor alphas ( $\alpha$ ), corresponding  $t$ -statistics ( $t_\alpha$ ), and  $R^2$ . The next two rows show spanning alphas ( $\alpha$ ), corresponding  $t$ -statistics ( $t_\alpha$ ), and  $R^2$  when regressing the long-short *ENS* returns on *OLS* returns and vice versa. All  $t$ -statistics are calculated using Newey and West (1987) adjusted standard errors with 4 lags. The sample period is from January 2002 to December 2021.

	Equal-Weighted					Value-Weighted				
	Avg	$t$	$\alpha$	$t_\alpha$	$R^2$	Avg	$t$	$\alpha$	$t_\alpha$	$R^2$
<i>OLS</i>	0.68	2.83	0.42	2.32	55.74	0.43	1.85	0.12	0.74	53.00
<i>ENS</i>	0.93	4.76	0.65	4.15	51.69	0.66	3.22	0.32	2.04	48.20
<i>ENS ~ OLS</i>			0.42	5.82	89.49			0.31	4.01	84.55
<i>OLS ~ ENS</i>			-0.43	-5.45	89.49			-0.26	-3.21	84.55

Table F5

**Limits to arbitrage: Summary statistics.** This table reports the summary statistics of limits-to-arbitrage proxies of different machine learning quintile portfolios. All stocks are sorted into country-neutral quintile portfolios based on their predicted returns for the next month. The sorting breakpoints are based on big stocks only, which are in the top 90% of the country's aggregated market capitalization. We compute for each quintile portfolio the average monthly value of each of three proxies for limits to arbitrage:  $-1 \times$  market capitalization (*SIZE*), idiosyncratic volatility (*IVOL*), Amihud illiquidity (*ILLIQ*), and a combination of the different proxies (*COMBO*). All proxies for limits to arbitrage are ranked into the  $[-1, 1]$  interval for each month and country and higher values indicate higher limits to arbitrage. Afterward, we report the time-series average. The sample period is from January 2002 to December 2021.

	<i>SIZE</i>		<i>IVOL</i>		<i>ILLIQ</i>		<i>COMBO</i>	
	<i>OLS</i>	<i>ENS</i>	<i>OLS</i>	<i>ENS</i>	<i>OLS</i>	<i>ENS</i>	<i>OLS</i>	<i>ENS</i>
Low (L)	0.10	-0.11	0.03	0.02	0.33	0.33	0.15	0.15
2	0.01	0.06	-0.02	-0.03	0.11	-0.06	0.03	-0.05
3	-0.03	0.10	-0.03	-0.05	-0.04	-0.14	-0.03	-0.10
4	-0.06	0.09	-0.03	-0.03	-0.17	-0.18	-0.09	-0.10
High (H)	-0.05	-0.04	0.03	0.04	-0.36	-0.13	-0.13	-0.02

Table F6

**Robustness - Further investment frictions.** This table reports the performance of different buy/hold long-only strategies before and after transaction costs. The investment universe is limited to big stocks. We investigate predictions from a linear OLS model and an ensemble (ENS) of non-linear machine learning models (RF, GBRT, and NN<sub>1-5</sub>). Every month the portfolio consists of the stocks that currently exhibit the highest X% forecasted returns per country plus those selected in previous months whose forecasted returns have not deteriorated beyond the top Y%. The first number in the column row names represents X, while the second represents Y. We report the strategies' gross returns in excess of the market, average two-way turnover, transaction costs, net returns in excess of the market, and net Fama and French (2018) six-factor models alphas. We assume one-way transaction costs of 100 basis points. All *t*-statistics are Newey and West (1987) adjusted with 4 lags. Panel A summarizes results from equal-weighting while Panel B shows results from value-weighting. The sample period is from January 2002 to December 2021.

	OLS		ENS	
	20%/20%	10%/30%	20%/20%	10%/30%
<b>Panel A: Equal-weighted</b>				
$r_{gross}^e - Mkt$	0.47 (5.31)	0.44 (5.07)	0.78 (7.88)	0.78 (7.37)
TO (in %)	44.14	24.68	45.13	27.38
T-cost (in %)	0.44	0.25	0.45	0.27
$r_{net}^e - Mkt$	0.03 (0.37)	0.19 (2.26)	0.33 (3.35)	0.51 (4.82)
$\alpha_{net}^{FF6}$	0.17 (2.77)	0.33 (5.17)	0.43 (5.84)	0.62 (7.35)
<b>Panel B: Value-weighted</b>				
$r_{gross}^e - Mkt$	0.27 (2.99)	0.29 (3.13)	0.45 (4.67)	0.46 (4.53)
TO (in %)	44.09	21.86	45.46	23.28
T-cost (in %)	0.44	0.22	0.45	0.23
$r_{net}^e - Mkt$	-0.17 (-1.88)	0.07 (0.74)	-0.01 (-0.08)	0.23 (2.26)
$\alpha_{net}^{FF6}$	-0.16 (-3.27)	0.06 (1.14)	0.02 (0.36)	0.25 (3.72)

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